## Development and validation of the HLS-EU-Q12

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## 1. Background

The European Health Literacy Questionnaire (HLS-EU-Q47) consists of 47 items, which according to the underlying conceptual model, address a matrix of 3 by 4 domains resulting in 12 elements of the health literacy (HL) conceptual matrix (cf. Sørensen et al., 2013; Sørensen et al., 2015; Pelikan \& Ganahl, 2017). Accordingly, the 47 items assess self-reported difficulties in the four cognitive domains accessing, understanding, appraising and applying information relevant for taking decisions in the three health domains healthcare, disease prevention and health promotion (Sørensen et al., 2013; Sørensen et al., 2015). Participants are asked to rate each item on a 4-point Likert like scale (very easy, fairly easy, fairly difficult, very difficult). Furthermore, they have the option to choose "don't know".

The items were developed in English and then translated into Bulgarian, Dutch, German, Greek, Polish and Spanish. The psychometric properties of the questionnaire was investigated using Principal Component Analysis (PCA) and reliability analysis using data from a field test conducted in Ireland and the Netherlands (for more details on the development process see Sørensen et al., 2013; Sørensen et al., 2015). The HLS-EU-Q47 was applied in the first wave of the European Health Literacy Survey in eight countries (HLS-EU-8): Austria (AT), Germany (only North-Rhine-Westphalia, DE), Spain (ES), Ireland (IE), The Netherlands (NL), Bulgaria (BG), Poland (PL), and Greece (EL). Data was collected either by Computer Assisted Personal Interviewing (CAPI) or Paper Assisted Personal Interviewing (PAPI). Recruitment strategies varied between countries (cf. Pelikan \& Ganahl, 2017).

Using data from HLS-EU-8, four main index scores were constructed for "general HL" (comprising all 47 items), "healthcare literacy", "disease prevention literacy" and "health promotion literacy", and reliability for these indexes was assessed using Cronbach's $\alpha$. The Cronbach $\alpha$ 's for all four indexes across all eight countries were at least 0.87 and the item correlations with the total scales exceeded 0.30 (HLS-EU Consortium, 2012). Furthermore, in order to justify the usage of an overall sum score, Item Response Theory (IRT) analysis was applied to examine unidimensionality of the HLS-EU-Q47. The Rating Scale Model (RSM; Andrich, 1978) was used with the four-point scale and the Rasch Model (RM) with dichotomized data (very easy / fairly easy vs. fairly difficult / very difficult). The RSM analysis showed poor model fit. To test the fit of the RM to the data Likelihood Ratio Tests (Andersen, 1973) using the split criteria median test score, gender and dichotomized educational level were conducted for each of the eight countries. As result from these analyses a 16 -item version was proposed (HLS-EUQ16; only unpublished manuscript available; for more details see Pelikan \& Ganahl, 2017). Correlations between the indexes of this short version and the 47-item version varied between $r=0.73$ and $r=0.88$ in the different countries (cf. Pelikan \& Ganahl, 2017). However, the HLS-EU-Q16 does not include an item of the element „apply information" in the "health promotion" domain of the HL conceptual matrix.

In the last years, both in Norway (HLS-Q12; Finbråten et al., 2017; Finbråten et al., 2018) and in Taiwan (HL-SF12; Duong et al., 2017) 12-item versions of the HLS-EU were developed in which each of the elements of the HL conceptual matrix is represented by one item. Whereas Finbråten et al. (2017) and Finbråten et al. (2018) applied Confirmatory Factor Analysis (CFA) and IRT, Duong et al. (2017) used only CFA to examine the psychometric properties of their 12-item version. However, only $50 \%$ of the items of these two 12-item versions are overlapping (see Table 1). Four of this six items are also contained in the HLS-EU-Q16.

Table 1: Items included in the HLS-EU-Q16, the HLS-Q12 (Finbråten et al., 2017; Finbråten et al., 2018) and the HL-SF12 (Duong et al., 2017)

| HL conceptual matrix element | Item number in the HLS-EU-Q47 |  |  |
| :--- | :---: | :---: | :---: |
|  | HLS-EU-Q16 | HLS-Q12 | HL-SF12 |
| 1 (acess information, healthcare) | 2,4 | 2 | 2 |
| 2 (understand information, healthcare) | 5,8 | 7 | 6 |
| 3 (appraise information, healthcare) | 11 | 10 | 10 |
| 4 (apply information, healthcare) | 13,16 | 14 | 15 |
| 5 (acess information, disease prevention) | 18 | 18 | 18 |
| 6 (understand information, disease prevention) | 21,23 | 23 | 23 |
| 7 (appraise information, disease prevention) | 28 | 28 | 26 |
| 8 (apply information, disease prevention) | 31 | 30 | 30 |
| 9 (acess information, health promotion) | 33 | 32 | 33 |
| 10 (understand information, health promotion) | 37,39 | 38 | 39 |
| 11 (appraise information, health promotion) | 43 | 43 | 43 |
| 12 (apply information, health promotion) | - | 44 | 45 |

A short version representing all 12 elements of the HL conceptual matrix by one item which sufficiently meets the requirements of a unidimensional IRT model is highly desirable for several reasons. Therefore, in preparation of the second wave of the European Health Literacy Survey ( $\mathrm{HLS}_{19}$ ) new IRTanalyses using data from HLS-EU-8 were conducted with the goal to select a subsample of items - the HLS-EU-Q12 - which should fulfill the following criteria.

The HLS-EU-Q12 should

1. represent all 12 elements of the HL conceptual matrix by one item,
2. include as many items from the HLS-EU-Q16 as possible (cf. Table 1),
3. show the greatest possible overlap with the HLS-Q12 (Finbråten et al., 2017; Finbråten et al., 2018; cf. Table 1), and
4. represent a close to optimal 12-item solution, i.e. the solution with the lowest deviance from the assumptions of the Partial Credit Model (PCM; Masters, 1982) when analyzed separately for each HLS-EU-8 country.

In the following, the development of the HLS-EU-Q12 based on HLS-EU-8 data and its validation using data from $\mathrm{HLS}_{19}$ is described.

## 2. Development of the HLS-EU-Q12

### 2.1 Methods

## Participants

Analyses are based on data from all eight countries of the HLS-EU-8 study collected in 2011. A detailed description of the HLS-EU-8 recruitment strategies in the different countries can be found elsewhere (e.g. Sørensen et al., 2015; Pelikan \& Ganahl, 2017). Across all HLS-EU-8 countries data from $\mathrm{n}=8102$ persons were available whereby sample sizes varied between $n=1000$ and $n=1057$ in the individual countries (see Table 2).

Table 2: Sample sizes in the 8 HLS-EU countries

|  | Sample size |  |
| :---: | :---: | ---: |
| Country | AT | 1015 |
|  | BG | 1002 |
|  | EL | 1000 |
|  | ES | 1000 |
|  | IE | 1005 |
|  | NL | 1023 |
|  | PL | 1000 |
|  | DE | 1057 |
| Total | 8102 |  |

## Data analysis

The data set was divided randomly into a training data set $(n=4054)$ and a test data set ( $n=4048$ ). An iterative IRT analysis approach combined with expert judgement on content validity was chosen, including the following steps:

1) PCM analysis of HLS-EU-Q47 on the training data set across all HLS-EU-8 countries ( $n=4054$ ): The goal was to find additional items on top of HLS-EU-Q16 which could be used for item selection for the HLS-EU-Q12.
2) Selection of additional items based on the results of the PCM analysis (Step 1) and expert judgement on content validity (exclude low priority items).
3) PCM analysis of HLS-EU-Q16 plus additional items chosen in Step 2 using the test data set ( $\mathrm{n}=$ 4048) for each HLS-EU-8 country separately: The aims of this step were to evaluate the item selection of Step 2 for each of the HLS-EU-8 country and to find the HLS-EU-Q12 solution with the best fit to the PCM.
4) PCM analysis of the selected 12 items (from Step 3) on the same test data set ( $\mathrm{n}=4048$ ) for each HLS-EU-8 country. Since in Step 3 some items have been removed, the remaining 12 items were retested to evaluate if the scale has been affected (cf. Robinson et al., 2019).
5) Comparison of PCM model fit of the different questionnaire versions (HLS-EU-Q47, HLS-EUQ16, HLS-EU-Q12, HLS-Q12, HL-SF12) using the test data set $(\mathrm{n}=4048)$ for each HLS-EU-8 country, examination of the correlations of the HLS-EU-Q12 with the HLS-EU-Q47, the HLS-EU-

Q16 and the Newest Vital Sign test (NVS; Weiss et al., 2005¹), and calculation of Cronbach's $\alpha$ as well as item-total correlations for the HLS-EU-Q12. In order to calculate the correlations of the HLS-EU-Q12 with the HLS-EU-Q47, HLS-EU-Q16 and the NVS, indices of HL were constructed as described in Sørensen et al. (2015).

## PCM analysis:

All analyses were conducted in R 3.5.1 (https://cran.r-project.org/) using the packages TAM 3.1-45 (Robitzsch, Kiefer \& Wu, 2019, 2020), sirt 3.3.-26 (Robitzsch, 2019), and mirt (Chalmers, 2012; version 1.30). Persons with more than 3 missing values were excluded. The PCM with ConQuest parametrization was used (Robitzsch, Kiefer \& Wu, 2019).

In Steps 1, 3 and 4, item infit statistics and corresponding $t$-statistics were calculated for each item. The expected value is 1 ; values $>1$ indicate that the item is less predictable than what would be expected according to the IRT model (underfit), values < 1 mean that the item is more predictable than what would be expected according to the expectations of the IRT model (= overfit; Linacre \& Wright, 1994, p. 360). Underfitting items may severely degrade the measurement, whereas overfitting items may overestimate raw score differences (Smith et al., 2008). The Holm procedure was applied to adjust the $p$-values for multiple testing (cf. Robitzsch, Kiefer \& Wu, 2019). Items were interpreted as over/underfitting if the adjusted $p$-value was $\leq 0.05$. The Nominal Categories Model was applied to check whether the expected ordering of response categories is supported by the data (Thissen, Cai \& Bock, 2010; Chalmers et al., 2019, p. 100). Differential item functioning (DIF) analyses were conducted using gender and the dichotomized criteria age (median split) and education (< higher education entrance qualification vs. at least higher education entrance qualification). A facets analysis was conducted. The criteria were set up as facets (e.g. for gender, item+gender+item*gender), and the IRT analysis was rerun (Robitzsch, Kiefer \& Wu, 2019). The interaction term item*gender yields the DIF magnitude.

For the comparison of PCM model fit of the different questionnaire versions (Step 5), SRMSR (standardized root mean square residual; Maydeu-Olivares, 2013) was calculated for each of the questionnaire versions (cf. Robitzsch, Kiefer \& Wu, 2019). SRMSR is a global fit statistic based on the comparison of residual correlations of item pairs. Maydeu-Olivares suggests a cutoff of $\leq 0.05$ for wellfitting IRT models. A less conservative value of 0.08 often is used as acceptable (cf. Hu \& Bentler, 1999). Furthermore, the combined PCA / t-test protocol to examine unidimensionality (cf. Smith, 2002; Hagell, 2014) was applied to the different versions. Two subsets of items are formed based on a PCA of standardized item residuals pursuant to the loadings of the item residuals on the first principal component (cf. Hagell, 2014). Person parameter estimation is conducted in each of the two item subsets and the resulting person parameter estimates from the two subsets are compared by means of paired $t$-tests (cf. Hagell, 2014). Under the assumption of unidimensionality, the proportion of individuals with significantly different person parameters in the two item subsets is small, i.e. $\leq 5 \%$ of the $t$-tests are significant, or the lower bound of a $95 \%$ confidence interval (CI) of the observed proportion overlaps 5\% (Hagell, 2014). In our analysis the Agresti-Coull CI was used. WLE reliability and EAP (expected a posteriory) reliability coefficients were calculated according to Adams (2005) (cf. Robitzsch, Kiefer \& Wu, 2019). Additionally, deviance, Akaike’s Information Criterion (AIC; Akaike, 1973), the AIC correction for small samples (AICc; Hurvich \& Tsai, 1989), Bozdogan's (1987) consistent

[^0]AIC (CAIC) and the Bayesian Information Criterion (BIC; Schwarz, 1978) were calculated to compare the data-model fit for the different versions. Lower values indicate better data-model fit.

### 2.2 Results

## Step 1:

No unordered response categories were observed. Seven items of the HLS-EU-Q16 had significant infit statistics (see Table A1 in the Appendix). Overfit was observed for six items (items 13, 21, 23, 33, 39, 43 ), and underfit was observed for item 28 with an infit statistic of 1.10 ( $t=4.60, p<0.001$ ). Another 17 items of the remaining items of the HSL-EU-Q47 had significant infit statistics. DIF was observed for 12 items of the HLS-EU-Q16 (items 5, 18, 39 for age; items $8,21,23,28,31,37$ for education, and items 2, 11 and 33 for age and education). For 6 items of the HLS-EU-Q47 which are not included in the HLS-EU-Q16 neither over-/underfit nor DIF was observed (see Table A1). As in previous analyses, the most problematic subdomain was "health promotion"; only for two items of this subdomain neither DIF nor over-/underfit was observed.

## Step 2:

Six items were candidates to be selected as additional items on top of the 16 items of the HLS-EU-Q16 according to the results of Step 1. Two of them were judged as low priority items and were not considered. Thus, four items were selected: 7, 10, 24, 44. Furthermore, it was decided to include two additional items from the health promotion domain, although they showed DIF in the training data set (items 36 and 42), such that each of the four cognitive domains (access, understand, appraise and apply) was represented by two items. This resulted in six additional items, whereby three of them are included in the HLS-Q12 (Finbråten et al., 2017).

## Step 3:

Using the test data set, only for items 28 (infit: 1.23, $t=3.47, p=0.032$ ) and 36 (infit: 1.46, $t=5.75, p$ < 0.001) significant underfit was observed in Germany (see Table A2 in the Appendix). DIF for age was observed for item 2 in four countries (AT, EL, ES, NL), for item 33 in two countries ( $B G, E L$ ), and for item 23, 39 and 42 in one country. DIF for gender was found for items 5 and 43 in one country, and DIF for education was found for item 8 in two countries as well as for items 11,21 and 31 in one country.

The proposed solution fulfilling the abovementioned criteria (represent all 12 elements of the HL conceptual matrix by one item, include as many items of the HLS-EU-Q16 as possible, show the greatest possible overlap with the HLS-Q12 and showing the lowest deviance from the assumptions of the PCM across all HLS-EU-8 countries) consisted of the items $4,7,10,16,18,23,24,31,33,37,42,44$.

## Step 4:

PCM analysis of the selected 12 items in each of the 8 countries revealed no significant infit statistics, however DIF for education for item 31 in Austria and for item 33 in Bulgaria, and DIF for age for item 33 in Bulgaria and Greece as well as for item 42 in The Netherlands (see Table A3 in the Appendix).

## Step 5:

SRMSR-values for the HLS-EU-Q12 were < 0.08 in the individual countries and thus are acceptable (Table 3). For the HLS-Q12 values $>0.08$ were observed in two countries and for the HL-SF12 in three countries.

Table 3: SRMSR values for the different versions in the eight countries

| Country | HLS-EU-Q47 | HLS-EU-Q16 | HLS-EU-Q12 | HLS-Q12 | HL-SF12 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| AT | $\mathbf{0 . 0 9 2 0}$ | $\mathbf{0 . 0 8 2 8}$ | 0.0712 | 0.0768 | 0.0775 |
| BG | $\mathbf{0 . 0 8 9 1}$ | 0.0773 | 0.0696 | 0.0713 | $\mathbf{0 . 0 9 0 0}$ |
| EL | $\mathbf{0 . 1 0 3 1}$ | $\mathbf{0 . 0 9 1 2}$ | 0.0769 | $\mathbf{0 . 0 9 3 6}$ | $\mathbf{0 . 1 0 3 1}$ |
| ES | $\mathbf{0 . 0 9 2 0}$ | 0.0748 | 0.0683 | 0.0653 | 0.0760 |
| IE | $\mathbf{0 . 0 9 4 9}$ | $\mathbf{0 . 0 8 8 5}$ | 0.0789 | 0.0776 | 0.0716 |
| NL | $\mathbf{0 . 0 9 4 7}$ | $\mathbf{0 . 0 8 8 5}$ | 0.0745 | 0.0742 | 0.0767 |
| PL | 0.0795 | 0.0643 | 0.0595 | 0.0543 | 0.0644 |
| DE | $\mathbf{0 . 1 1 2 9}$ | $\mathbf{0 . 0 9 2 1}$ | 0.0798 | $\mathbf{0 . 1 1 0 5}$ | $\mathbf{0 . 0 8 9 8}$ |

For each of the three 12 -item versions the proportions of significant $t$-tests were $>5 \%$ in all countries, and only in one country the lower bound of the $95 \% \mathrm{Cl}$ included $5 \%$ for each of the versions (Table 4). In three countries the proportion of significant $t$-tests was lowest for the HLS-EU-Q12 (ES, IE, NL), and in one country for the HL-SF12 (BG). In AT, EL and DE the proportion was comparable for the HLS-EUQ12 and HL-SF12, and in PL it was comparable for the HLS-Q12 and HL-SF12.

Table 4: Results of PCA/t-test procedure (proportion of significant t-tests and lower bound of CI)

| Country | HLS-EU-Q47 | HLS-EU-Q16 | HLS-EU-Q12 | HLS-Q12 | HL-SF12 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| AT | $0.259(0.224)$ | $0.184(0.154)$ | $0.086(0.065)$ | $0.121(0.096)$ | $0.082(0.061)$ |
| BG | $0.310(0.271)$ | $0.158(0.129)$ | $0.125(0.099)$ | $0.104(0.080)$ | $0.087(0.066)$ |
| EL | $0.271(0.234)$ | $0.159(0.130)$ | $0.101(0.077)$ | $0.114(0.089)$ | $0.098(0.076)$ |
| ES | $0.253(0.217)$ | $0.179(0.147)$ | $0.101(0.077)$ | $0.113(0.088)$ | $0.143(0.115)$ |
| IE | $0.289(0.251)$ | $0.139(0.111)$ | $0.084(0.062)$ | $0.102(0.078)$ | $0.104(0.080)$ |
| NL | $0.216(0.182)$ | $0.141(0.112)$ | $0.069(0.049)$ | $0.107(0.082)$ | $0.074(0.054)$ |
| PL | $0.234(0.199)$ | $0.103(0.079)$ | $0.078(0.057)$ | $0.066(0.047)$ | $0.063(0.044)$ |
| DE | $0.291(0.253)$ | $0.146(0.118)$ | $0.078(0.058)$ | $0.143(0.115)$ | $0.076(0.056)$ |

When comparing the three 12 -item versions by means of deviance and information criteria, the HLS-EU-Q12 showed best fit to the PCM (i.e. consistently had the lowest values in seven of the eight countries); in Austria the HL-SF12 had the lowest values (see Table A4 in the Appendix). All three 12item versions had acceptable WLE and EAP reliability coefficients $>0.77$ in all eight countries (see Table 5).

Table 5: WLE and EAP reliability coefficients for the different questionnaire versions

| Country | HLS-EU-Q47 | HLS-EU-Q16 | HLS-EU-Q12 | HLS-Q12 | HL-SF12 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| WLE reliability coeff. |  |  |  |  |  |
| AT | 0.948 | 0.868 | 0.830 | 0.828 | 0.835 |
| BG | 0.965 | 0.912 | 0.884 | 0.884 | 0.882 |
| EL | 0.952 | 0.886 | 0.846 | 0.851 | 0.850 |
| ES | 0.951 | 0.873 | 0.830 | 0.837 | 0.827 |
| IE | 0.945 | 0.874 | 0.838 | 0.839 | 0.839 |
| NL | 0.938 | 0.839 | 0.781 | 0.784 | 0.771 |
| PL | 0.962 | 0.901 | 0.870 | 0.881 | 0.873 |
| DE | 0.946 | 0.878 | 0.829 | 0.827 | 0.820 |

EAP reliability coeff.

| AT | 0.955 | 0.877 | 0.840 | 0.836 | 0.845 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| BG | 0.972 | 0.927 | 0.902 | 0.897 | 0.897 |
| EL | 0.965 | 0.906 | 0.867 | 0.872 | 0.871 |
| ES | 0.954 | 0.876 | 0.835 | 0.841 | 0.829 |
| IE | 0.960 | 0.897 | 0.868 | 0.864 | 0.865 |
| NL | 0.944 | 0.861 | 0.813 | 0.807 | 0.800 |
| PL | 0.970 | 0.917 | 0.893 | 0.898 | 0.889 |
| DE | 0.958 | 0.905 | 0.860 | 0.850 | 0.850 |

The correlation of the HLS-EU-Q12 and the HLS-EU-Q47 indices was high in the total sample of all eight countries ( $r=0.957$ ). In the individual countries the correlations varied between 0.938 and 0.967 (see Table 6). The correlations with the HLS-EU-Q16 were comparable. The correlation of the HLS-EU-Q12 index with the NVS was $r=0.26$ in the total sample and the correlations in the individual countries varied between $r=0.13$ and $r=0.269$. These values are comparable to the correlations of the HLS-EUQ47 index with the NVS ( $r=0.25$ for the total EU-8, and correlations between $r=0.14$ and $r=0.38$ in the individual countries; cf. Pelikan \& Ganahl, 2017).

Table 6: Correlations of the HLS-EU-Q12 with HLS-EU-Q16 and NVS

|  | AT | BG | EL | ES | IE | NL | PL | DE | Total (EU-8) |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| HLS-EU-Q47 | 0,946 | 0,967 | 0,960 | 0,938 | 0,963 | 0,938 | 0,962 | 0,961 | 0,957 |
| HLS-EU-Q16 | 0,930 | 0,970 | 0,952 | 0,931 | 0,953 | 0,929 | 0,965 | 0,945 | 0,951 |
| NVS | 0,153 | 0,399 | 0,322 | 0,210 | 0,269 | 0,190 | 0,392 | 0,130 | 0,26 |

## Replacing item 33 by item 32 and evaluating model fit

Following a consortium decision, it was examined if item 33 could be replaced by item 32 . Therefore, Steps 4 and 5 were applied to a version consisting of items $4,7,10,16,18,23,24,31,32,37,42,44$ (called HLS-EU-Q12 ${ }_{32}$ in the following) in order to evaluate its model fit.

For the HLS-EU-Q12 ${ }_{32}$ no significant infit statistics were observed (see Table A5 in the Appendix), as was the case for the version containing item 33 instead of item 32. DIF for age was observed for item 32 in two countries (BG, EL) and DIF for education in two countries (BG, IE). In three countries SRSMRvalues $>0.08$ were observed for the HLS-EU-Q1232, and the proportion of significant t-tests was $<5 \%$ only in one country (see Table 7). WLE und EAP reliability coefficients were comparable for both test versions with values $>0.77$ for HLS-EU-Q12 ${ }_{32}$ and $>0.78$ for the HLS-EU-Q12 in all countries. Comparing the two versions by deviance and information statistics, the version containing item 32 shows consistently lower values across all statistics and across all countries.

Table 7: Comparison of HLS-EU-Q12 und HLS-EU-Q1232

| AT | BG | EL | ES | IE | NL | PL | DE |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |

## SRMSR

| HLS-EU-Q12 | 0.0712 | 0.0696 | 0.0769 | 0.0683 | 0.0789 | 0.0745 | 0.0595 | 0.0798 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| HLS-EU-Q1232 | 0.0728 | 0.0699 | $\mathbf{0 . 0 8 0 6}$ | 0.0693 | $\mathbf{0 . 0 8 5 2}$ | 0.0739 | 0.0631 | 0.0881 |

PCA/t-test (proportion significant t-tests, Cl )

| HLS-EU-Q12 | $\mathbf{0 . 0 8 6}$ | $\mathbf{0 . 1 2 5}$ | $\mathbf{0 . 1 0 1}$ | $\mathbf{0 . 1 0 1}$ | $\mathbf{0 . 0 8 4}$ | 0.069 | $\mathbf{0 . 0 7 8}$ | $\mathbf{0 . 0 7 8}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | $\mathbf{( 0 . 0 6 5 )}$ | $\mathbf{( 0 . 0 9 9 )}$ | $\mathbf{( 0 . 0 7 7 )}$ | $\mathbf{( 0 . 0 0 7 )}$ | $\mathbf{( 0 . 0 6 2 )}$ | $(0.049)$ | $\mathbf{( 0 . 0 5 7 )}$ | $\mathbf{( 0 . 0 5 8 )}$ |
| HLS-EU-Q1232 | $\mathbf{0 . 0 9 5}$ | $\mathbf{0 . 0 7 9}$ | $\mathbf{0 . 1 2 2}$ | $\mathbf{0 . 0 7 9}$ | $\mathbf{0 . 1 0 4}$ | 0.039 | $\mathbf{0 . 0 8 6}$ | $\mathbf{0 . 1 0 4}$ |
|  | $(0.073)$ | $\mathbf{( 0 . 0 5 8 )}$ | $\mathbf{( 0 . 0 9 6 )}$ | $\mathbf{( 0 . 0 5 8 )}$ | $\mathbf{( 0 . 0 8 )}$ | $(0.025)$ | $\mathbf{( 0 . 0 6 4 )}$ | $\mathbf{( 0 . 0 8 )}$ |

## WLE reliability coeff

| HLS-EU-Q12 | 0.8304 | 0.8838 | 0.8464 | 0.8298 | 0.8383 | 0.7806 | 0.8703 | 0.8293 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| HLS-EU-Q1232 | 0.8271 | 0.8842 | 0.8444 | 0.8315 | 0.8357 | 0.7761 | 0.8710 | 0.8308 |
| EAP reliability coff. |  |  |  |  |  |  |  |  |
| HLS-EU-Q12 | 0.8304 | 0.8838 | 0.8464 | 0.8298 | 0.8383 | 0.8137 | 0.8933 | 0.8601 |
| HLS-EU-Q1232 | 0.8368 | 0.9010 | 0.8654 | 0.8365 | 0.8666 | 0.8112 | 0.8939 | 0.8622 |

## Deviance

| HLS-EU-Q12 | 13130.16 | 11309.64 | 11713.98 | 9370.28 | 10756.19 | 10749.74 | 9835.49 | 12001.58 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| HLS-EU-Q1232 | 13074.86 | 11294.01 | 11694.90 | 9269.33 | 10570.27 | 10516.50 | 9792.51 | 11895.35 |

AIC

| HLS-EU-Q12 | 13204.16 | 11383.28 | 11787.98 | 9444.56 | 10830.19 | 10823.74 | 9909.49 | 12075.58 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| HLS-EU-Q1232 | 13148.86 | 11368.01 | 11768.90 | 9343.33 | 10644.27 | 10590.50 | 9866.51 | 11969.35 |
| AICC |  |  |  |  |  |  |  |  |
| HLS-EU-Q12 | 13209.79 | 11389.25 | 11793.86 | 9450.79 | 10836.39 | 10829.85 | 9915.67 | 12081.51 |
| HLS-EU-Q1232 | 13154.49 | 11373.98 | 11774.77 | 9349.55 | 10650.47 | 10596.61 | 9872.69 | 11975.28 |

CAIC

| HLS-EU-Q12 | 13399.74 | 11576.88 | 11982.16 | 9636.75 | 11022.54 | 11016.53 | 10101.91 | 12269.40 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| HLS-EU-Q1232 | 13344.44 | 11561.61 | 11963.08 | 9535.52 | 10836.62 | 10783.29 | 10058.93 | 12163.17 |

## BIC

| HLS-EU-Q12 | 13362.74 | 11539.88 | 11945.16 | 9599.75 | 10985.54 | 10979.53 | 10064.91 | 12232.40 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| HLS-EU-Q1232 | 13307.44 | 11524.61 | 11926.08 | 9498.52 | 10799.62 | 10746.29 | 10021.93 | 12126.17 |

The correlation of the HLS-EU-Q12 32 and the HLS-EU-Q47 indices was $r=0.955$ in the total sample of all eight countries and therefore comparable with the correlation of the version containing item 33 . In the individual countries the correlations varied between $r=0.935$ and $r=0.966$. The correlations with the HLS-EU-Q16 were also comparable. The correlation of the HLS-EU-Q12 index with the NVS was $r=$ 0.263 in the total sample and the correlations in the individual countries varied between $r=0.13$ and $r=0.385$. Therefore, all correlations are comparable with the version containing item 33 instead of item 32.

Table 8: Correlations of the HLS-EU-Q1232 with HLS-EU-Q16 and NVS

|  | AT | BG | EL | ES | IE | NL | PL | DE | Total (EU-8) |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| HLS-EU-Q47 | 0.949 | 0.966 | 0.959 | 0.938 | 0.954 | 0.935 | 0.956 | 0.959 | 0.955 |
| HLS-EU-Q16 | 0.929 | 0.969 | 0.945 | 0.925 | 0.942 | 0.919 | 0.959 | 0.940 | 0.946 |
| NVS | 0.168 | 0.406 | 0.314 | 0.211 | 0.266 | 0.195 | 0.385 | 0.130 | 0.263 |

## 3. The HLS ${ }_{19}-\mathrm{Q} 12$

According to a consortium decision, in $\mathrm{HLS}_{19}$ the 12 -item version containing items $4,7,10,16,18,23$, $24,31,32,37,42,44$, with some improvement in the wording of a few items, and a change in the wording of the response categories (omitting the qualifier "fairly" in the two middle categories, resulting in the four point scale "very easy", "easy", "difficult", "very difficult") were used and named $\mathrm{HLS}_{19}-\mathrm{Q} 12$. Figure 1 shows the instruction and items of the $\mathrm{HLS}_{19}-\mathrm{Q} 12$.

Figure 1: Instruction and items of the HLS ${ }_{19}-Q 12$

| INTRODUCTION | Interviewer: It is not always easy to get understandable, reliable and useful information on health related to <br> With the following questions we would like to find out which tasks related to handling health information ar difficult. |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| CORE-HL | On a scale from very easy to very difficult, how easy would you say it is: |  |  |  |
| [SHOWCARD WITH SCALE - ONE ANSWER PER ROW] |  |  |  |  |
|  |  | Very easy | Easy | Difficult |


| CORE-HL4 | ...to find out where to get professional help when you are ill? <br> [instructions: such as doctor, nurse, pharmacist, psychologist] |
| :---: | :--- |
| CORE-HL7 | ...to understand information about what to do in a medical emergency? |
| CORE-HL10 | ...to judge the advantages and disadvantages of different treatment options? |
| CORE-HL16 | ...to act on advice from your doctor or pharmacist? |
| CORE-HL18 | ...to find information on how to handle mental health problems? <br> [Instruction: stress, depression or anxiety] |
| CORE-HL23 | ...to understand information about recommended health screenings or <br> examinations? <br> [Instructions: e.g. colorectal cancer screening, blood sugar test] |
| CORE-HL24 | ...to judge if information on unhealthy habits, such as smoking, low physical <br> activity or drinking too much alcohol, are reliable? |
| CORE-HL31 | ...to decide how you can protect yourself from illness using information from the <br> mass media? <br> [Instructions: e.g. Newspapers, TV or Internet] |
| CORE-HL32 | ...to find information on healthy life styles such as physical exercise, healthy food <br> or nutrition? |
| CORE-HL37 | ...to understand advice concerning your health from family or friends? <br> CORE-HL42...to judge how your housing conditions may affect your health and well-being? <br> CORE-HL44...to make decisions to improve your health and well-being? |

## 4. Validation of the HLS $\mathrm{H}_{19}-\mathrm{Q} 12$

The aim was to evaluate the PCM model fit of the $\mathrm{HLS}_{19}-\mathrm{Q} 12$ using the data of the second HLS-EU wave. Furthermore, it was aimed to test the model fit of the dichotomized version to the Rasch model.

### 3.1 Methods

## Participants

Analyses are based on data from 15 countries of the second wave of the HLS-EU study collected between November 2019 and February 2021. A detailed description of the recruitment strategies in the different countries can be found in Chapter 2 of the $\mathrm{HLS}_{19}$ International Report.

Table 9 gives an overview on the data collection methods, sample sizes and version of the questionnaire used (all 47 items of the HLS-EU-Q47, the HLS-Q16 plus 6 items on top selected in step 2 of the $\mathrm{HLS}_{19}-\mathrm{Q} 12$ development, or only the 12 items of the $\mathrm{HLS}_{19}-\mathrm{Q} 12$ ) for the individual countries. Across all countries data from $n=38080$ persons were available whereby sample sizes varied between $n=1000$ and $n=5660$ in the individual countries.

Table 9: Overview on analyzed data for the validation of the HLS 19 -Q12

| Country | Data collection method | HLS version | N | $\mathrm{N} \quad$ without missing items |
| :---: | :---: | :---: | :---: | :---: |
| Austria (AT) | CATI | HLS ${ }_{19}$-Q12 | 2967 | 2471 |
| Belgium (BE) | CAWI | HLS-Q16 plus 6 | 1000 | 1000 |
| Czech Republic (CZ) | CATI, CAWI | HLS-Q16 plus 6 | 1599 | 1459 |
| Denmark (DK) | CAWI | HLS-Q16 plus 6 | 3602 | 3506 |
| Germany (DE) | PAPI | HLS-EU-Q47 | 2143 | 1991 |
| France (FR) | CAWI | HLS-Q16 plus 6 | 2003 | 2003 |
| Hungary | CATI | HLS-Q16 plus 6 | 1195 | 1021 |
| Ireland (IE) | CATI | HLS-EU-Q47 | 4487 | 4142 |
| Israel (IL) | CATI, CAWI | HLS-Q16 plus 6 | 1315 | 1294 |
| Norway (NO) | CATI | HLS-EU-Q47 | 2855 | 2387 |
| Portugal (PT) | CAWI | HLS ${ }_{19}$-Q12 | 1247 | 922 |
| Russia (RU) | PAPI | HLS-Q16 plus 6 | 5660 | 4752 |
| Slovenia (SI) | CAPI, Paper, CAWI | HLS-EU-Q47 | 3360 | 3178 |
| Slovakia (SK) | CAPI | HLS-Q16 plus 6 | 2145 | 2144 |
| Switzerland (CH) | CATI, CAWI | HLS ${ }_{19}$-Q12 | 2502 | 2370 |

## Data analysis

The assumption of unidimensionality for $\mathrm{HLS}_{19}-\mathrm{Q} 12$ was tested by means of the PCM in a first step and the dichotomous Rasch Model (RM) in a subsequent step. All analyses were conducted in $R$ using the packages eRm 1.0-1 (Mair et al., 2018), TAM 3.5-19 (Robitzsch, Kiefer \& Wu, 2020) and mirt 1.33.2 (Chalmers, 2015). Persons with missing values on at least one of the items were excluded. Analyses were conducted for each of the countries separately. Due to very large sample sizes in some countries (e.g. RU and IE), all analyses were conducted also in a random sample of $n=900$ for each of the countries (the sample size for the random sample $n=900$ ), and PCM analyses on item level were also conducted in four randomly chosen independent subsamples in each of the countries (therefore the sample sizes in the four subsamples vary according to the total sample sizes in the individual countries). Due to the huge number of significance tests and large sample sizes, $\alpha=0.001$ was chosen.

For PCM analysis the same methods were applied as in the development of the $\mathrm{HLS}_{19}-\mathrm{Q} 12$ described in 2.1:

- Individual items:
- item infit statistics and corresponding $t$-statistics were calculated for each item;
- the Nominal Categories Model was applied to check whether the expected ordering of response categories is supported by the data;
- DIF analyses were conducted using the split criteria gender, median age and education (< higher education entrance qualification vs. at least higher education entrance qualification)
- Global fit statistics:
- SRMSR (standardized root mean square residual;
- combined PCA / t-test protocol;
- WLE reliability and EAP (expected a posteriori) reliability coefficients.

Additionally, local stochastic independence was assessed by means of an adjusted variant of the Q3 statistic by Yen (1984), aQ3, for all item pairs and an effect size of model fit (MADaQ3), which is the average of the absolute values of $a Q 3$ statistics and $p$-values adjusted according to the Holm procedure (Robitzsch, Kiefer \& Wu, 2020). For those countries, in which different data collection methods were applied (CH, CZ, IL, SI), analyses were conducted both independent of survey-type as well as separately for the survey-types. However, sample sizes were very small for CATI especially in CH ( $n=139$ ). Therefore, no additional analyses were performed with random samples in survey-type-specific analyses.

For the dichotomous scoring, Likelihood-Ratio-Tests (LR-test, Andersen, 1973) were conducted as global model tests using median test score, education (< higher education entrance qualification vs. at least higher education entrance qualification), median age and gender as split criteria, and individual item-fit statistics (Fischer-Scheiblechner z-statistics, Fischer \& Scheiblechner, 1970) were calculated. A global test for local independence, which calculates the sum of absolute deviations between the observed inter-item correlations and the expected correlations (Mair et al., 2020), was conducted. Furthermore, on item level increased correlations between inter-item residuals were checked by means of the Q3-statistic (cf. Mair et al., 2020). RM analyses were conducted in the total sample and the random sample of $n=900$. Furthermore, graphical model tests according to Rasch (1960/1980) were applied to examine model fit for each individual item. These model tests rely on the assumption that item parameters can be consistently estimated in different subsamples drawn from a population in which the Rasch model applies. Additionally, item characteristic curve (ICC) plots were used to graphically inspect model fit of the individual items. The ICC plots show how the probability for response category 1 expected by the RM changes with the values of the latent variable. Observed
scores are represented by circles. If the deviations of the observed values from the expected values are small, there is close conformity of the data with the model.

### 3.2 Results

## PCM analysis

Global fit statistics:
The WLE and EAP reliability coefficients have acceptable values $>0.78$ in all countries and all survey types (cf. Table 10). For those countries which were included in the development of the HLS ${ }_{19}-\mathrm{Q} 12$, the coefficients are comparable in AT, however lower in IE and DE (cf. Table 7, HLS-EU-Q12 ${ }_{32}$ ). Furthermore, Cronbach $\alpha$ 's are sufficient in all countries and all survey types with values $>0.80$.

Table 10: WLE and EAP reliability coefficients of the HLS ${ }_{19}$-Q12

| Country | WLE rel. | EAP rel. | Cronbach $\alpha$ | N |
| :---: | :---: | :---: | :---: | :---: |
| AT | 0.830 | 0.850 | 0.843 | 2471 |
| BE | 0.879 | 0.883 | 0.881 | 1000 |
| CH_Total | 0.839 | 0.844 | 0.836 | 2370 |
| CH_CAWI | 0.841 | 0.847 | 0.838 | 2231 |
| CH_CATI | 0.811 | 0.808 | 0.801 | 139 |
| CZ_Total | 0.845 | 0.850 | 0.842 | 1459 |
| CZ_CAWI | 0.848 | 0.852 | 0.844 | 1057 |
| CZ_CATI | 0.820 | 0.823 | 0.814 | 402 |
| DE | 0.811 | 0.809 | 0.802 | 1991 |
| DK | 0.853 | 0.862 | 0.857 | 3506 |
| FR | 0.878 | 0.894 | 0.887 | 2003 |
| HU | 0.841 | 0.845 | 0.843 | 1021 |
| IE | 0.785 | 0.822 | 0.822 | 4142 |
| IL_Total | 0.874 | 0.890 | 0.882 | 1294 |
| IL_CAWI | 0.869 | 0.879 | 0.871 | 1004 |
| IL_CATI | 0.855 | 0.889 | 0.887 | 290 |
| NO | 0.831 | 0.844 | 0.838 | 2387 |
| PT | 0.839 | 0.853 | 0.902 | 922 |
| RU | 0.884 | 0.892 | 0.900 | 4752 |


| Country | WLE rel. | EAP rel. | Cronbach $\alpha$ | N |
| :--- | ---: | ---: | ---: | ---: |
| SI_Total | 0.873 | 0.887 | 0.892 | 3178 |
| SI_CAWI | 0.847 | 0.864 | 0.864 | 1463 |
| SI_CAPI | 0.887 | 0.899 | 0.912 | 1704 |
| SK | 0.881 | 0.887 | 0.884 | 2144 |

According to the PCA/t-test procedure, the $\mathrm{HLS}_{19}-\mathrm{Q} 12$ cannot be considered sufficiently unidimensional except for NO, as the proportion of individuals with significant different person parameters in two item subsets exceeds $5 \%$ in all countries except for Norway in the random sample, and only for Norway the lower bound of the confidence interval (Cl) overlaps $5 \%$ also in the total sample (see Table 11). In IE the lower bound of the Cl is $<0.06$ both in the total and the random sample, in IL in the total sample and in RU in the random sample, and in FR and CZ the lower bound of the Cl is $<0.07$ in the total sample. The highest percentages are observable in HU, PT and SK both in the total and the random samples. For SI results for the two survey modes are comparable; regarding the results for CATI in CH and IL it has to be considered that the samples sizes are small ( $n=139$ and $n=290$, respectively). For those countries which were included in the development of the $\mathrm{HLS}_{19}-\mathrm{Q} 12$ (AT, IE, DE), the results are slightly better or comparable than in the previous analysis (cf. Table 7, HLS-EUQ12 ${ }_{32}$ ).

Table 11: Results of SRMSR and PCA/t-test procedure for the HLS $19-Q 12$

| Country | SRMSR | PCA/t-test |  |
| :--- | :--- | :--- | :--- |
|  |  | Total sample | Random sample |
| AT | 0.0596 | $0.093(0.082)$ | $0.077(0.061)$ |
| BE | 0.0659 | $0.088(0.072)$ | $0.093(0.076)$ |
| CH_Total | 0.0641 | $0.089(0.078)$ | $0.111(0.092)$ |
| CH_CAWI | 0.0651 | $0.096(0.084)$ | - |
| CH_CATI | 0.1104 | $0.165(0.112)$ | - |
| CZ_Total | 0.0541 | $0.080(0.067)$ | $0.104(0.086)$ |
| CZ_CAWI | 0.0559 | $0.092(0.076)$ | - |
| CZ_CATI | 0.0777 | $0.117(0.089)$ | - |
| DE | 0.0679 | $0.091(0.080)$ | $0.110(0.091)$ |
| DK | 0.0566 | $0.112(0.102)$ | $0.108(0.089)$ |
| FR | 0.0591 | $0.074(0.063)$ | $0.086(0.069)$ |
| HU | 0.0781 | $0.146(0.126)$ | $0.141(0.120)$ |
| IE | 0.0699 | $0.061(0.054)$ | $0.069(0.054)$ |


| IL_Total | 0.0485 | $0.070(0.057)$ | $0.083(0.070)$ |
| :--- | :--- | :--- | :--- |
| IL_CAWI | 0.0538 | $0.088(0.072)$ | - |
| IL_CATI | 0.0789 | $0.072(0.047)$ | - |
| NO | 0.0632 | $0.057(0.049)$ | $0.021(0.013)$ |
| PT | 0.0753 | $0.140(0.119)$ | $0.124(0.104)$ |
| RU | 0.0504 | $0.090(0.082)$ | $0.068(0.053)$ |
| SI_Total | 0.0775 | $0.085(0.075)$ | $0.113(0.094)$ |
| SI_CAWI | 0.0784 | $0.090(0.076)$ | - |
| SI_CAPI | 0.0672 | $0.090(0.078)$ | - |
| SK | 0.0556 | $0.116(0.103)$ | $0.128(0.107)$ |

The SRMSR statistics are above the cut-off value of 0.05 for good model fit suggested by MaydeuOlivares (2013) in all countries except IL (see Table 11). However, the values are below the less conservative cut-off value of 0.08 according to $\mathrm{Hu} \&$ Bentler (1999). But the global test for local independence based on the adjusted Q3-statistic yielded significant results in all countries both in the total and the random samples (see Table 12). For CATI the Q3-statistic was not significant in CH and IL; however, the sample sizes are very small.

Table 12: Results of the model tests for local independence based on the adjusted Q3-statistic in all countries in the total and the random samples

|  | Total Sample |  |  | Random Sample |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Country | MADaQ3 | max. aQ3 | $\boldsymbol{p}$ | MADaQ3 | max. aQ3 | $\boldsymbol{p}$ |
| AT | 0.061 | 0.255 | $<0.001$ | 0.058 | 0.246 | $<0.001$ |
| BE | 0.070 | 0.200 | $<0.001$ | 0.073 | 0.200 | $<0.001$ |
| CH_Total | 0.061 | 0.287 | $<0.001$ | 0.073 | 0.274 | $<0.001$ |
| CH_CAWI | 0.061 | 0.286 | $<0.001$ | - | - | - |
| CH_CATI | 0.108 | 0.309 | 0.012 | - | - | - |
| CZ_Total | 0.045 | 0.183 | $<0.001$ | 0.050 | 0.187 | $<0.001$ |
| CZ_CAWI | 0.051 | 0.223 | $<0.001$ | - | - | - |
| CZ_CATI | 0.060 | 0.209 | 0.001 | - | - | - |
| DE | 0.067 | 0.170 | $<0.001$ | 0.069 | 0.201 | $<0.001$ |
| DK | 0.060 | 0.160 | $<0.001$ | 0.062 | 0.172 | $<0.001$ |
| FR | 0.061 | 0.266 | $<0.001$ | 0.059 | 0.300 | $<0.001$ |
| HU | 0.071 | 0.252 | $<0.001$ | 0.070 | 0.235 | $<0.001$ |


|  | Total Sample |  |  | Random Sample |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Country | MADaQ3 | max. aQ3 | $\boldsymbol{p}$ | MADaQ3 | max. aQ3 | $\boldsymbol{p}$ |
| IE | 0.061 | 0.249 | $<0.001$ | 0.071 | 0.282 | $<0.001$ |
| IL_Total | 0.060 | 0.200 | $<0.001$ | 0.063 | 0.221 | $<0.001$ |
| IL_CAWI | 0.066 | 0.200 | $<0.001$ | - | - | - |
| IL_CATI | 0.071 | 0.201 | 0.036 | - | - | - |
| NO | 0.053 | 0.157 | $<0.001$ | 0.056 | 0.183 | $<0.001$ |
| PT | 0.078 | 0.299 | $<0.001$ | 0.077 | 0.304 | $<0.001$ |
| RU | 0.068 | 0.220 | $<0.001$ | 0.066 | 0.244 | $<0.001$ |
| SI_Total | 0.059 | 0.244 | $<0.001$ | 0.067 | 0.306 | $<0.001$ |
| SI_CAWI | 0.052 | 0.212 | $<0.001$ | - | - | - |
| SI_CAPI | 0.066 | 0.258 | $<0.001$ | - | - | - |
| SK | 0.064 | 0.221 | $<0.001$ | 0.072 | 0.284 | $<0.001$ |

Analyses at item level:
The Q3-statistics for the item pairs in the random samples showed that the residuals of 14 item pairs are significantly correlated in several countries (see Table 13, Figure 2: Matrix of dependent item pairs and domains), also if the different survey modes are considered separately. Non significant results for CATI in $\mathrm{CH}, \mathrm{CZ}$ and IL are partly due to the small sample sizes.

Table 13: Dependent item pairs in the different countries according to the random samples

| Item pair | Countries |
| :---: | :---: |
| 4 (access, HC) - 7 (understand, HC) | AT, $\mathrm{CH}^{\text {g }}, \mathrm{CZ}^{\mathrm{h}}, \mathrm{DE}, \mathrm{DK}, \mathrm{NO}, \mathrm{PT}^{\text {a,d }}, \mathrm{SI}^{\mathrm{h}}, \mathrm{SK}^{\mathrm{c}}$ |
| 10 (appraise, HC) - 32 (access, HP) | ```AT}\mp@subsup{}{}{\textrm{b},\textrm{e}},\mp@subsup{\textrm{BE}}{}{\textrm{e}},\mp@subsup{\textrm{CH}}{}{\textrm{b},f,h},C\mp@subsup{Z}{}{c,f,g},D\mp@subsup{E}{}{f},D\mp@subsup{K}{}{c,f}, IL c,f,g SK``` |
| 23 (understand, DP) - 24 (appraise, DP) | DE, $\mathrm{FR}^{\text {a,f }}, \mathrm{HU}^{\text {d }}, I E, I L^{\mathrm{g}}, \mathrm{RU}^{\mathrm{b}}, \mathrm{SI}^{\text {a,e,h }}$ |
| 7 (understand, HC) - 10 (appraise, HC) | BE, $\mathrm{CH}^{\mathrm{g}}, \mathrm{FR}, \mathrm{IL}^{\mathrm{g}}, \mathrm{RU}^{\text {c }}$, SK |
| 24 (appraise, DP) - 32 (access, HP) | AT, CH ${ }^{\mathrm{g}}, \mathrm{CZ}^{\mathrm{g}}, \mathrm{DK}, \mathrm{PT}^{\text {a,d }}, \mathrm{Sl}^{\text {h }}$ |
| 42 (appraise, HP) - 44 (apply, HP) | AT, CH ${ }^{\text {b,f,g, }}, \mathrm{IE}^{\text {c }}, \mathrm{RU}, \mathrm{SK}$ |
| 16 (apply, HC) - 18 (access, DP) | $\mathrm{BE}^{\text {c,f }}, \mathrm{CZ}{ }^{\text {g }}$, PT |
| 16 (apply, HC) - 31 (apply, DP) | $\mathrm{BE}^{\text {c,f }}$, DE, PT |
| 32 (access, HP) - 37 (understand, HP) | $\mathrm{IL}^{\mathrm{h}}, \mathrm{PT}^{\text {a,d }}, \mathrm{SK}$ |
| 4 (access, HC) - 31 (apply, DP) | BE ${ }^{\text {b,e }}, \mathrm{DE}^{\text {c }}$ |
| 4 (access, HC) - 42 (appraise, HP) | $\mathrm{BE}^{\text {c,f }}, \mathrm{DE}^{\text {c,f }}$ |
| 16 (apply, HC) - 32 (access, HP) | $\mathrm{BE}, \mathrm{IL}^{\mathrm{g}}, \mathrm{PT}^{\mathrm{b}, \mathrm{e}}$ |
| 37 (understand, HP) - 42 (appraise, HP) | $\mathrm{IE}, \mathrm{SK}, \mathrm{PT}^{\text {d }}$ |

HC: healthcare; DP: disease prevention; HP: health promotion; ${ }^{a}$ : $r>0.30$ in the random sample; ${ }^{b}: r>0.25$ in the random sample; ${ }^{c}: r>0.20$ in the random sample; ${ }^{d}: r>0.30$ in the total sample; ${ }^{e}: r>0.25$ in the total sample; ${ }^{f}$ : $r>0.20$ in the total sample; ${ }^{\text {g }}$ : total sample and CAWI; ${ }^{\text {n }}$ : in all survey modes;

Two dependent item pairs were observed in the healthcare ( HC ) domain, one in the disease prevention (DP) domain and three in the health promotion (HP) domain in several countries and with a correlation of the residuals of $r>0.20$ in at least one country in the total and/or the random sample. In the cognitive domain "apply information" one dependent item pair was observed. The remaining six dependent item pairs appeared across the domains.

Some more dependent item pairs were found only in PT or SK with correlations $r>0.20$, with some correlations in PT above 0.30 .

Figure 2: Matrix of dependent item pairs and domains

|  |  | HC |  |  | DP |  |  |  | HP |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Items | 7 | 10 | 16 | 18 | 23 | 24 | 31 | 32 | 37 | 42 | 44 |
| HC | 4 | AT, CH, CZ, DE, DK, NO, PT, SI, SK |  |  |  |  |  | BE, DE |  | PT | $\begin{aligned} & \hline \mathrm{BE}, \\ & \mathrm{DE} \end{aligned}$ |  |
|  | 7 |  | $\mathrm{BE}, \mathrm{CH}$, <br> FR, IL, <br> RU, SK |  |  |  |  |  |  |  | SK | SK |
|  | 10 |  |  |  |  |  |  |  | AT, BE, CH, CZ, DE, DK, IL, SK |  |  |  |
|  | 16 |  |  |  | $\begin{aligned} & \text { BE, CZ, } \\ & \text { PT } \end{aligned}$ |  |  | $\begin{aligned} & \mathrm{BE}, \mathrm{DE}, \\ & \text { PT } \end{aligned}$ | BE, IL, PT | PT |  |  |
| DP | 23 |  |  |  |  |  | DE, FR, HU, IE, IL, RU, SI | SI |  |  |  |  |
|  | 24 |  |  |  |  |  |  |  | $\begin{aligned} & \text { AT ,CH, CZ, } \\ & \text { DK, PT, SI } \\ & \hline \end{aligned}$ | PT |  |  |
|  | 31 |  |  |  |  |  |  |  |  |  |  |  |
| HP | 32 |  |  |  |  |  |  |  |  | PT, SK | PT |  |
|  | 37 |  |  |  |  |  |  |  |  |  | $\begin{aligned} & \hline \mathrm{IE}, \\ & \mathrm{PT}, \\ & \mathrm{SK} \\ & \hline \end{aligned}$ |  |
|  | 42 |  |  |  |  |  |  |  |  |  |  | AT, CH, IE, RU, SK |

Table 14 shows the results for the item infit statistics, for the DIF analyses (only significant results are included in the table), and for the Nominal Response Model to identify the empirical ordering of the response categories. For Item 31 values of the infit statistic between 1.15 and 1.35 , with high values of the corresponding $t$-statistics and $p$-values $<0.001$, were observed in the total samples in four countries ( $\mathrm{CH}, \mathrm{IE}, \mathrm{NO}, \mathrm{SI}$ ). In SI (infit statistic $=1.36, t=6.99, p<0.001$ ) Item 31 displays also significant underfit in the random sample of $n=900$ and in all four independent subsamples as well as for both survey types. In IE (infit statistic $=1.19, t=4.23, p=0.001$ ) and NO (infit statistic $=1.18, t=3.76, p=0.006$ ) the infit statistics are also high in the random sample of $n=900$, and in CH for CAWI (infit statistic 1.15, $t=5.07, p<0.001$ ). Furthermore, Item 4 shows significant underfit in FR in the total sample and the random sample of $n=900$. Item 23 displays underfit in IE, Item 37 in DE in the total sample, and Item 44 in SI in the total sample and the CAPI sample. All other significant infit statistics indicate overfit of
the respective items which is, however, less problematic (for the infit statistics of all items in the total samples of the different countries see Table A6 in the Appendix).

Eleven of the items display DIF in at least one country for one split criterion in the total samples, and several items display DIF in more than one country also in the random sample of $n=900$ and/or at least one of the four independent subsamples (see

Table 14):

- Item 4 for
- gender in three countries (AT, DE, IE), and
- age in one country (AT)
- Item 7 for
- age in two countries (DK, IE), and
- education in IE,
- Item 10 for
- education in three countries ( $\mathrm{AT}, \mathrm{DK}, \mathrm{SI}^{2}$ )
- Item 16 for
- gender in IE
- Item 18 for
- age in IE, and
- education in RU
- Item 23 for
- gender in three countries ( $\mathrm{DK}, \mathrm{IE}, \mathrm{SI}{ }^{3}$ ),
- age in six countries ( $\mathrm{AT}, \mathrm{BE}, \mathrm{CH}^{4}, \mathrm{DK}, \mathrm{FR}, \mathrm{SI}^{3}$ ) and
- education in two countries ( $\mathrm{DK}, \mathrm{SI}^{2}$ ),
- Item 24 for
- age in IE, underfit SI
- Item 31 for
- gender in one country $\left(\mathrm{SI}^{4}\right)$,
- education in five countries ( $\mathrm{AT}, \mathrm{CZ}^{2}, \mathrm{IE}, \mathrm{NO}, \mathrm{SI}^{4}$ ), and
- age in one two countries ( $\mathrm{BE}, \mathrm{SI}^{2}$ )
- Item 32 for
- gender in one country (DK),
- age in three countries $\left(C Z^{3}, R U, S I^{3}\right)$ and
- education in five countries ( $\mathrm{AT}, \mathrm{CZ}^{4}, \mathrm{FR}, \mathrm{SI}{ }^{3}, \mathrm{SK}$ ),
- Item 37 for
- age in four countries ( $\mathrm{AT}, \mathrm{CH}^{4}, \mathrm{DK}, \mathrm{RU}$ ),
- education in DE, and
- Item 42 for
- age in four countries ( $\left.\mathrm{CH}^{2}, \mathrm{DK}, \mathrm{IE}, \mathrm{NO}\right)$, and
- education in IE.

[^1]Across countries, the most problematic items are items 4, 23, 31, 32, 37 and 42 both in the total and the random samples, and also if the different survey types are analyzed separately in $\mathrm{CH}, \mathrm{CZ}, \mathrm{IL}$ and $\mathrm{SI}{ }^{5}$.

Table 14: Results for infit statistics and DIF-analyses for the HLS ${ }_{19}-Q 12$

| Country | Item | Infit | $t$ | $p$ | Gender DIF |  | Age DIF |  | Education DIF |  | unordered response categories |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Magn. | $p$ | Magn. | $p$ | Magn. | $p$ |  |
| AT | 4 | 1.029 | 0.93 | 1.000 | 0.15 | <0.001 ${ }^{\text {a,b }}$ | 0.13 | <0.001 |  |  |  |
|  | 10 | 1.068 | 2.46 | 0.334 |  |  |  |  | 0.11 | $<0.001^{\text {a }}$ |  |
|  | 23 | 0.903 | -3.35 | 0.023 |  |  | 0.10 | <0.001 |  |  |  |
|  | 31 | 1.088 | 3.03 | 0.067 |  |  |  |  | 0.13 | $<0.001^{\text {b }}$ |  |
|  | 32 | 0.949 | -1.77 | 1.000 |  |  |  |  | 0.14 | $<0.001^{\text {b }}$ | x |
|  | 37 | 1.063 | 2.17 | 0.607 |  |  | 0.10 | <0.001 |  |  |  |
|  | 42 | 0.976 | -0.81 | 1.000 |  |  |  |  |  |  | $\mathrm{X}^{\text {a }}$ |
| BE | 23 | 0.904 | -2.31 | 0.612 |  |  | 0.21 | <0.001 ${ }^{\text {a }}$ |  |  |  |
|  | 31 | 1.177 | 3.94 | 0.003 |  |  | 0.17 | <0.001 |  |  |  |
| CH_Total | 16 | 0.973 | -0.79 | 1.000 |  |  |  |  |  |  | x |
|  | 23 | 0.975 | -0.81 | 1.000 |  |  | 0.14 | <0.001 ${ }^{\text {b }}$ |  |  | Random ${ }^{\text {f }}$ |
|  | 31 | 1.150 | 5.13 | <0.001 |  |  |  |  |  |  |  |
|  | 37 | 1.108 | 3.34 | 0.027 |  |  | 0.18 | <0.001 ${ }^{\text {d }}$ |  |  |  |
|  | 42 | 0.971 | -0.98 | 1.000 |  |  | 0.10 | <0.001 |  |  |  |
| CH_CAWI | 16 |  |  |  |  |  |  |  |  |  | x |
|  | 23 |  |  |  |  |  | 0.14 | <0.001 |  |  |  |
|  | 31 | 1.153 | 5.07 | <0.001 |  |  |  |  |  |  |  |
|  | 37 |  |  |  |  |  | 0.17 | <0.001 |  |  |  |
| CH_CATI | 10 |  |  |  |  |  |  |  | 0.42 | <0.001 | x |
|  | 16 |  |  |  |  |  |  |  |  |  | x |
| $\square$ 18 |  |  |  |  |  |  |  |  |  |  | X |
|  | 23 |  |  |  |  |  |  |  |  |  | x |
|  | 24 |  |  |  |  |  |  |  |  |  | X |

[^2]| Country | Item | Infit | $t$ | $p$ | Gender DIF |  | Age DIF |  | Education DIF |  | unordered response categories |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Magn. | $p$ | Magn. | $p$ | Magn. | $p$ |  |
|  | 31 |  |  |  |  |  |  |  |  |  | X |
|  | 32 |  |  |  |  |  |  |  |  |  | X |
|  | 37 |  |  |  |  |  |  |  |  |  | X |
|  | 44 |  |  |  |  |  |  |  |  |  | X |
| CZ_Total | 31 | 0.965 | -0.92 | 1.000 |  |  |  |  | 0.12 | <0.001 |  |
|  | 32 | 1.052 | 1.29 | 1.000 |  |  | 0.21 | $<0.001^{\mathrm{a}},$ <br> b | 0.20 | $<0.001{ }^{\text {b }}$ |  |
| CZ_CAWI | 32 |  |  |  |  |  | 0.17 | <0.001 | 0.18 | <0.001 |  |
| CZ_CATI | 4 |  |  |  |  |  |  |  |  |  | X |
|  | 7 |  |  |  |  |  |  |  |  |  | X |
|  | 10 |  |  |  |  |  |  |  |  |  | X |
|  | 16 |  |  |  |  |  |  |  |  |  | X |
|  | 18 |  |  |  |  |  |  |  |  |  | X |
|  | 23 |  |  |  |  |  |  |  |  |  | X |
|  | 24 |  |  |  |  |  |  |  |  |  | X |
|  | 31 |  |  |  |  |  |  |  |  |  | X |
|  | 32 |  |  |  |  |  | 0.39 | <0.001 | 0.29 | <0.001 | X |
|  | 37 |  |  |  |  |  |  |  |  |  | X |
|  | 42 |  |  |  |  |  |  |  | 0.23 | <0.001 | X |
|  | 44 |  |  |  |  |  |  |  |  |  | X |
| DE | 4 | 1.026 | 0.80 | 1.000 | 0.10 | <0.001 |  |  |  |  | $X^{\text {a }}$ |
|  | 23 | 0.940 | $-1.88$ | 1.000 |  |  |  |  |  |  | $X^{\text {a }}$ |
|  | 37 | 1.150 | 4.38 | <0.001 |  |  |  |  | 0.11 | <0.001 ${ }^{\text {a }}$ | $X^{\text {a }}$ |
| DK | 7 | 1.036 | 1.46 | 1.000 |  |  | 0.09 | <0.001 |  |  |  |
|  | 10 | 0.963 | -1.60 | 1.000 |  |  |  |  | 0.07 | <0.001 |  |
|  | 23 | 0.975 | -1.04 | 1.000 | 0.11 | $<0.001{ }^{\text {b }}$ | 0.17 | $<0.001^{\mathrm{a}},$ | 0.10 | <0.001 ${ }^{\text {b }}$ |  |
|  | 32 | 0.896 | -4.45 | <0.001 | 0.12 | $<0.001{ }^{\text {d }}$ |  |  |  |  |  |
|  | 37 | 1.073 | 2.86 | 0.103 |  |  | 0.09 | <0.001 |  |  |  |




| Country | Item | Infit | $t$ | $p$ | Gender DIF |  | Age DIF |  | Education DIF |  | unordered response categories |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Magn. | $p$ | Magn. | $p$ | Magn. | $p$ |  |
|  | 32 | 0.954 | -1.90 | 0.860 |  |  | 0.06 | <0.001 |  |  |  |
|  | 37 | 1.077 | 2.87 | 0.091 |  |  | 0.08 | $<0.001^{\text {b }}$ |  |  |  |
| SI_Total | 10 | 1.073 | 2.76 | 0.070 |  |  |  |  | 0.10 | <0.001 |  |
|  | 23 | 0.995 | -0.19 | 1.000 | 0.14 | $<0.001^{\text {a }}$ | 0.21 | $<0.001^{\text {a }}$ <br> e | 0.14 | <0.001 ${ }^{\text {a, }}$ |  |
|  | 24 | 0.840 | -6.16 | <0.001 |  |  |  |  |  |  |  |
|  | 31 | 1.350 | 12.69 | $<0.001^{\mathrm{a}},$ e | 0.09 | <0.001 | 0.10 | <0.001 | 0.21 | $<0.001^{a},$ e |  |
|  | 32 | 0.893 | -3.98 | 0.001 |  |  | 0.23 | $<0.001^{\text {a }}$ | 0.28 | <0.001 ${ }^{\text {a }}$ |  |
|  | 44 | 1.134 | 4.88 | <0.001 |  |  |  |  |  |  |  |
| SI_CAWI | 7 |  |  |  |  |  | 0.14 | <0001 |  |  |  |
|  | 23 |  |  |  | 0.17 | <0.001 | 0.45 | <0.001 |  |  |  |
|  | 31 | 1.328 | 8.34 | <0.001 | 0.12 | <0.001 |  |  | 0.17 | <. 001 |  |
|  | 32 |  |  |  |  |  | 0.16 | <0.001 | 0.22 | <0.001 |  |
|  | 37 |  |  |  |  |  |  |  |  |  | X |
|  | 42 |  |  |  |  |  |  |  |  |  | X |
|  | 44 | 1.181 | 4.67 | <0.001 |  |  |  |  |  |  |  |
| SI_CAPI | 23 |  |  |  | 0.12 | <0.001 | 0.18 | <0.001 |  |  |  |
|  | 24 | 0.825 | -4.65 | <0.001 |  |  |  |  |  |  |  |
|  | 31 | 1.275 | 7.14 | <0.001 |  |  |  |  |  |  |  |
|  | 32 |  |  |  |  |  | 0.17 | <0.001 | 0.20 | <0.001 |  |
| SK | 32 | 0.886 | -3.63 | 0.010 |  |  |  |  | 0.10 | <0.001 ${ }^{\text {e }}$ |  |

Only significant results are shown in the table; a: significant also in the random sample of $n=900$; b: significant also in one of the four subsamples; c: significant also in two of the four subsamples; d: significant also in three of the four subsamples; e: significant also in all four subsamples; f: only in the random sample.

Considering the random sample, results are more favorable in most countries. In RU and SK none of the items displays significant DIF or significant misfit in the random sample.

Applying the Nominal Categories Model to identify the empirical ordering of the response categories revealed unordered response categories for all items in PT, for nine of the twelve items in IE, for five items in FR and for two or three items in AT, DE, HU, IL, and NO (for parameter estimates of the response categories see Table A7 in the Appendix). Closer inspection showed that the response
category "very difficult" was very rarely chosen (< $1 \%$ of persons) in some items in several countries (see frequency distributions for the response categories in the different countries in Figure A1 in the Appendix). It seems that it was hard for the respondents to discriminate between the response options "very difficult" and "difficult". On the other hand, in PT at least two third of the persons have chosen the answer 'easy' in all items (e.g. for Item 37 approximately $84 \%$, see Figure 3 ). In the case of low endorsement rates in some of the categories the estimation of the parameters of the Nominal Categories Model can be affected such that response categories are tagged as unordered (cf. GarcíaPérez, 2018). Therefore, regarding the results for the survey mode CATI in CH, CZ and IL, the small sample sizes must be considered.

Figure 3: Frequency distributions for the response categories in PT


X0: very difficult; X1: difficult; X2: easy; X3: very easy

## Dichotomous Rasch model:

## Global model tests:

The global LR-test using the split criterion median score was not possible for AT, PT and SI, because the data were not well-conditioned (Fischer, 1981). Well-conditioned data means that in every possible partition of the items into two non-empty item subsets at least one person has chosen the answer category 1 on one item in the first subset and answer category 0 on one item in the other subset. This is a necessary and sufficient condition for a unique solution of the conditional maximum likelihood estimates of the item parameters (Fischer, 1981). Of the remaining twelve countries, four LR-tests were significant in the total samples and two of them also in the random sample ( $p<0.001$; see

Table 15).

Furthermore, the global LR-tests using the split criterion median age was significant in 11 countries in the total samples, and for 7 countries also in the random sample. In RU the data in the random sample were not well-conditioned, thus the LR-test could not be performed. Using the split criterion gender the LR tests were significant in 3 countries only in the total sample. For the split criterion education the LR tests were significant in 9 countries in the total sample only. Only for CZ and HU none of the LR tests was significant in the total samples and random samples, and in the random sample none of the LR tests was significant for AT, IL, SI and SK.

Table 15: Results of the $L R$-tests of the dichotomized $H L S_{19}-Q 12$

| Split Criterion | Country | Total Sample |  |  | Random Sample |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Median Score |  | Chisq | DF | P | Chisq | DF | P |
|  | BE | 16.67 | 11 | 0.118 | 17.30 | 11 | 0.099 |
|  | CH | 45.63 | 11 | <0.001 | 15.64 | 11 | 0.155 |
|  | CZ | 18.65 | 11 | 0.068 | 12.08 | 11 | 0.358 |
|  | DE | 49.98 | 10 | <0.001 | 40.78 | 10 | <0.001 |
|  | DK | 32.72 | 11 | 0.001 | 8.67 | 11 | 0.652 |
|  | FR | 29.12 | 11 | 0.002 | 27.36 | 11 | 0.004 |
|  | HU | 20.41 | 11 | 0.040 | 18.38 | 11 | 0.073 |
|  | IE | 57.72 | 11 | <0.001 | 35.83 | 11 | <0.001 |
|  | IL | 21.08 | 11 | 0.033 | 19.54 | 11 | 0.052 |
|  | NO | 24.12 | 11 | 0.012 | 11.10 | 11 | 0.435 |
|  | RU | 84.40 | 11 | <0.001 | Not well-conditioned |  |  |
|  | SK | 29.78 | 11 | 0.002 | 15.28 | 11 | 0.17 |
| Age |  |  |  |  |  |  |  |
|  | AT | 38.74 | 11 | <0.001 | 23.05 | 11 | 0.017 |
|  | BE | 87.33 | 11 | <0.001 | 69.41 | 11 | <0.001 |
|  | CH | 159.84 | 11 | <0.001 | 50.96 | 11 | <0.001 |
|  | CZ | 28.12 | 11 | 0.003 | 23.30 | 11 | 0.016 |
|  | DE | 33.18 | 11 | <0.001 | 18.29 | 12 | 0.075 |
|  | DK | 176.77 | 11 | <0.001 | 51.78 | 11 | <0.001 |
|  | FR | 76.42 | 11 | <0.001 | 33.72 | 11 | <0.001 |


| Split Criterion | Country | Total Sample |  |  | Random Sample |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | HU | 15.83 | 11 | 0.148 | 13.81 | 11 | 0.244 |
|  | IE | 115.56 | 11 | <0.001 | 26.73 | 11 | 0.005 |
|  | IL | 20.66 | 11 | 0.037 | 20.17 | 11 | 0.043 |
|  | NO | 71.59 | 11 | <0.001 | 38.86 | 11 | <0.001 |
|  | PT | 43.22 | 11 | <0.001 | 44.67 | 11 | <0.001 |
|  | RU | 124.56 | 11 | <0.001 | 35.79 | 11 | <0.001 |
|  | SI | 99.50 | 11 | <0.001 | 23.88 | 11 | 0.013 |
|  | SK | 31.06 | 11 | 0.001 | 12.18 | 11 | 0.350 |
| Gender |  |  |  |  |  |  |  |
|  | AT | 28.55 | 11 | 0.003 | 19.56 | 11 | 0.052 |
|  | BE | 21.02 | 11 | 0.033 | 19.20 | 11 | 0.058 |
|  | CH | 43.06 | 11 | <0.001 | 16.96 | 11 | 0.109 |
|  | CZ | 28.45 | 11 | 0.003 | 10.65 | 11 | 0.473 |
|  | DE | 25.33 | 11 | 0.008 | 12.71 | 11 | 0.313 |
|  | DK | 68.07 | 11 | <0.001 | 21.55 | 11 | 0.028 |
|  | FR | 27.85 | 11 | 0.003 | 20.10 | 11 | 0.044 |
|  | HU | 10.84 | 11 | 0.457 | 8.21 | 11 | 0.695 |
|  | IE | 31.74 | 11 | 0.001 | 20.93 | 11 | 0.034 |
|  | IL | 18.20 | 11 | 0.077 | 16.92 | 11 | 0.110 |
|  | NO | 16.40 | 11 | 0.127 | 24.97 | 11 | 0.009 |
|  | PT | 17.30 | 11 | 0.099 | 15.61 | 11 | 0.156 |
|  | RU | 21.79 | 11 | 0.026 | 11.96 | 11 | 0.367 |
|  | SI | 35.56 | 11 | <0.001 | 16.17 | 11 | 0.135 |
|  | SK | 15.99 | 11 | 0.141 | 7.58 | 11 | 0.750 |
| Education |  |  |  |  |  |  |  |
|  | AT | 39.12 | 11 | <0.001 | 18.86 | 11 | 0.064 |
|  | BE | 9.29 | 11 | 0.595 | 8.81 | 11 | 0.64 |
|  | CH | 42.11 | 11 | <0.001 | 28.03 | 11 | 0.003 |
|  | CZ | 20.97 | 11 | 0.034 | 21.18 | 11 | 0.032 |


| Split Criterion | Country | Total Sample |  | Random Sample |  |  |  |
| :--- | :--- | ---: | ---: | ---: | :---: | :---: | :---: |
|  | DE | 40.47 | 11 | $<0.001$ | 29.61 | 11 | 0.002 |
|  | DK | 44.07 | 11 | $<0.001$ | 23.01 | 11 | 0.018 |
|  | FR | 17.39 | 11 | 0.097 | 15.50 | 11 | 0.161 |
|  | HU | 22.74 | 11 | 0.019 | 22.30 | 11 | 0.022 |
|  | IL | 32.08 | 11 | 0.001 | 13.75 | 11 | 0.247 |
|  | NO | 38.38 | 11 | $<0.001$ | 27.68 | 11 | 0.004 |
|  | PT | 27.83 | 11 | 0.003 | 26.27 | 11 | 0.006 |
|  | RU | 69.19 | 11 | $<0.001$ | 11.10 | 11 | 0.435 |
|  | SI | 78.07 | 11 | $<0.001$ | 27.68 | 11 | 0.004 |
|  | SK | 57.17 | 11 | $<0.001$ | 16.48 | 11 | 0.124 |

Only for AT and CZ the global model tests for local independence were not significant in the random samples (see Table 16: Results of the global model test for local independence in the total and the random samples).

Table 16: Results of the global model test for local independence in the total and the random samples

| Country | $\boldsymbol{p}$ total sample | $\boldsymbol{p}$ random sample |
| :--- | :--- | :--- |
| AT | $<0.001$ | 0.014 |
| BE | $<0.001$ | $<0.001$ |
| CH | $<0.001$ | $<0.001$ |
| CZ | $<0.001$ | 0.002 |
| DE | $<0.001$ | $<0.001$ |
| DK | $<0.001$ | $<0.001$ |
| FR | $<0.001$ | $<0.001$ |
| HU | $<0.001$ | $<0.001$ |
| IE | n.a. | $<0.001$ |
| IL | $<0.001$ | $<0.001$ |
| NO | $<0.001$ | $<0.001$ |
| PT | $<0.001$ | $<0.001$ |
| RU | n.a. | $<0.001$ |
| SI | $<0.001$ | $<0.001$ |
| SK | $<0.001$ | $<0.001$ |

n.a.: statistic could not be calculated due to the large sample size

## Analyses at item level:

Item 31 again has infit values $\geq 1.10$ with corresponding high $t$-values in several countries (CH, FR, NO, SI; cf. Table 17). Inspection of the ICC plots, for instance for NO and SI, reveals clear deviations of the observed scores from the expected values (cf. Figure 4).

NO


SI


The Fischer-Scheiblechner $z$-statistic (Wald test) was significant for several items and split criteria, therefore indicating DIF (cf. Table 17):

- Item 4 for
- gender in AT, IE and
- median test score in FR,
- Item 7 for age in IE, PT, RU,
- Item 10 for
- gender in DK,
- age in CH, DK, and
- education in AT, IL,
- Item 16 for
- age in RU,
- Item 18 for
- gender in DK,
- age in DK, IE, RU, and
- education in RU,
- Item 23 for
- gender in $\mathrm{CH}, \mathrm{DK}, \mathrm{SI}$
- age in $\mathrm{BE}, \mathrm{CH}, \mathrm{DE}, \mathrm{DK}, \mathrm{FR}, \mathrm{SI}$ and
- education in DK,
- Item 24 for median score in RU,
- Item 31 for
- education in AT, NO, SI and
- age in BE, DK, FR, NO
- median test score in CH ,
- Item 32 for
- gender in DK,
- age in NO, PT, RU, SI, and
- education in IL, RU, SI,
- Item 37 for
- gender in BE
- age in AT, CH, DK, NO, RU
- education in $\mathrm{CH}, \mathrm{DE}, \mathrm{SK}$, and
- median test score in DE,
- Item 42 for
- age in $\mathrm{CH}, \mathrm{DK}, \mathrm{IE}, \mathrm{NO}$ and
- Item 44 for education in SI.

Table 17: Results of item infit statistics and DIF analyses for the dichotomized $H L S_{19}-Q 12$

| Country | Item | Infit |  | Median DIF |  | Gender DIF |  | Age DIF |  | Education DIF |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | MSQ | t | z | P | z | P | z | P | z | P |
| AT | 4 | 0.90 | -1.89 |  |  | 4.22 | <0.001 |  |  |  |  |
|  | 10 | 1.03 | 1.60 |  |  |  |  |  |  | 3.57 | $<0.001^{\text {a }}$ |
|  | 31 | 0.97 | -1.11 |  |  |  |  |  |  | 3.40 | 0.001 |
|  | 37 | 1.05 | 1.51 |  |  |  |  | 4.44 | <0.001 |  |  |
| BE | 23 | 0.94 | -1.55 |  |  |  |  | -4.80 | <0.001 ${ }^{\text {a }}$ |  |  |
|  | 31 | 1.08 | 2.33 |  |  |  |  | 5.06 | $<0.001^{\text {a }}$ |  |  |
|  | 37 | 1.04 | 0.87 |  |  | -3.56 | <0.001 |  |  |  |  |
| CH | 10 | 1.01 | 0.25 |  |  |  |  | 3.63 | <0.001 |  |  |
|  | 23 | 0.97 | -0.95 |  |  | -3.61 | <0.001 | -6.44 | <0.001 ${ }^{\text {a }}$ |  |  |
|  | 31 | 1.11 | 4.99 | 4.88 | <0.001 |  |  |  |  |  |  |
|  | 37 | 1.09 | 2.71 |  |  |  |  | 7.97 | $<0.001^{\text {a }}$ | -4.30 | <0.001 |
|  | 42 | 0.95 | -1.73 |  |  |  |  | -4.41 | <0.001 |  |  |
| DE | 23 | 0.91 | -2.70 |  |  |  |  | -4.28 | <0.001 |  |  |
|  | 37 | 1.10 | 2.80 | 5.48 | <0.001 ${ }^{\text {a }}$ |  |  |  |  |  |  |
| DK | 10 | 0.94 | -3.32 |  |  | 3.51 | <0.001 | $4.73{ }^{\text {n }}$ | <0.001 | 3.25 | 0.001 |
|  | 18 | 0.97 | -1.69 |  |  | -3.88 | <0.001 | 4.38 | <0.001 |  |  |
| - 23 |  | 0.89 | -3.28 |  |  | -3.36 | 0.001 | -4.82 | <0.001 | -3.91 | <0.001 |
| - 31 |  | 1.01 | 0.47 |  |  |  |  | 3.72 | <0.001 |  |  |
| - 32 |  | 0.83 | -3.38 |  |  | -3.52 | <0.001 |  |  |  |  |
|  | 37 | 1.06 | 2.43 |  |  |  |  | 4.71 | <0.001 |  |  |
|  | 42 | 0.96 | -1.84 |  |  |  |  | -8.58 | <0.001 |  |  |
|  | 44 | 1.07 | 3.48 | 4.93 | <0.001 |  |  |  |  |  |  |
| FR | 4 | 1.08 | 2.00 | 4.03 | <0.001 ${ }^{\text {a }}$ |  |  |  |  |  |  |
|  | 23 | 0.92 | -2.33 |  |  |  |  | -5.58 | <0.001 |  |  |
|  | 31 | 0.97 | -1.16 |  |  |  |  | 4.71 | <0.001 |  |  |
| IE | 4 | 0.99 | -0.46 |  |  | 4.03 | <0.001 |  |  |  |  |


${ }^{\text {a. also significant in the random sample of } n=900 ; ~}{ }^{n}$ : non-uniform DIF

The check of local independence using the Q3-statistic revealed dependent item pairs in all countries except for AT in the random sample (cf. Table 18).

Table 18: Number of item pairs with significant correlations of inter-item residuals in the different countries

| Country | Total sample | Random sample |
| :--- | ---: | ---: |
| AT | 5 | 0 |
| BE | 6 | 5 |
| CH | 5 | 1 |
| CZ | 6 | 2 |
| DE | 5 | 3 |
| DK | 10 | 3 |
| FR | 5 | 4 |
| HU | n.a. | 4 |
| IE | 2 | 2 |
| IL | 6 | 2 |
| NO | 2 | 2 |
| PT | n.a. | 2 |
| RU | 4 | 2 |
| SI | 5 | 2 |
| SK | 5 | 2 |

## 5. Summary

The short version of the HLS-EU questionnaire representing all 12 elements of the HL conceptual matrix ( $\mathrm{HLS}_{19}$-Q12) which was developed using data from the first wave of the HLS-EU study did not show acceptable fit to the PCM in all 15 countries. According to the PCA/t-test procedure, only in Norway the $\mathrm{HLS}_{19}-\mathrm{Q} 12$ could be deemed sufficiently unidimensional. The values of the SRMSR statistic were above the cut-off value of 0.05 for good model fit suggested by Maydeu-Olivares (2013) in all countries except Israel, although below the less conservative value according to Hu and Bentler (1999). However, the global model test for local independence based on the adjusted Q3-statistic (Robitzsch et al., 2020) yielded significant results in all 15 countries both in the total samples as well as the random samples of $n=900$ drawn in each of the countries. Analyses on the item level revealed that for Russia and Slovakia none of the items displayed misfit, DIF or unordered response categories in the random samples of $n=900$. However, also in Russia and Slovakia - as in all other countries - some dependent item pairs were observed with residual correlations between $r=0.20$ and $r=0.25$ in the total or the random samples. In all other countries there was at least one item with unordered response categories, and/or at least one poor fitting item, and/or at least one item displaying DIF in the random sample, the total sample, and also in smaller subsamples. The most problematic items are $4,23,31$, 32,37 and 42, which showed significant model deviations in several countries. Items 18, 23, 31, 32, 37 and 42 already displayed DIF in previous analyses using the original HLS-EU-8 data (see 2.2 ). However, it needs to be considered that the response categories and also the wording of some items have been changed which of course could affect model fit either positively or negatively. The low endorsement rates in some of the categories could be the reason for the high number of items which were tagged as unordered in some countries (cf. García-Pérez, 2018). The frequency distribution for answer category "very difficult" varies from $0.30 \%$ (item 16) to $13.80 \%$ (item 7), for "difficult" from $3.15 \%$ (item 16) to $57.31 \%$ (item 10), for "easy" from $25.06 \%$ (item 10) to $83.95 \%$ (item 37), and for "very easy" from $3.82 \%$ (item 10) to $67.89 \%$ (item 16). However, the response category "very difficult" was very rarely chosen (< $1 \%$ of persons) in some items in several countries. Survey-type-specific analyses for CH, CZ, IL and SI revealed comparable results for CAWI and also for CAPI in SI. For CATI the results are somewhat different; however, sample sizes are (too) small, especially in CH ( $n=139$ ) and IL ( $n=290$ ).

For the dichotomous RM, model fit is barely acceptable for the Czech Republic and Austria. In both countries, in the random sample of $n=900$, neither the test of local dependency nor the LR tests were significant. However, in Austria item 10 displayed DIF for education both in the total and the random sample, and in the Czech Republic there are two dependent item pairs according to the Q3 statistic. In all other countries the global test for local independence and / or at least one LR test was significant both in the total and the random samples. On item level, for the split criteria gender, age and education the results for the dichotomous scoring are quite similar to those for the polytomous scoring. Again, items 23, 31, 32, 37 and 42 display DIF in several countries and/or split criteria.

Exploring the most problematic items leads to following hypotheses:

- Item 23 displays DIF for age in many countries, whereby the item is relatively easier for the older age group. The examples given in the instructions could be the reason.
- Item 31 is generally a poor fitting item and displays DIF for education and age in several countries. This item is easier for people of the lower education group and for the younger age group. The reason could be that higher educated people and older people are more critical regarding information from the mass media. Therefore, a second dimension of critical appraisal of information provided by mass media might be responsible for the poor model fit of this item.

In those four countries, in which different survey methods were applied, analyses were conducted across the survey methods. Additional separate analyses for the different survey methods could be conducted. However, the overall conclusion that some items display significant misfit to both the PCM and the RM in several countries would not be affected.

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Appendix

Table A1: Results of Step 1 of the development of the HLS-EU-Q12 in the total sample of the EU-8 countries

| Item | Infit | $t$ | p_Holm | DIF [magnitude] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Gender | Age | Education |
| 1 | 1.027 | 1.175 | 1.000 |  | o > y [0.196] | I > h [0.19] |
| 2 | 0.983 | -0.74 | 1.000 |  | o > y [0.107] | $1>\mathrm{h}$ [0.112] |
| 3 | 1.082 | 3.53 | 0.028 |  |  |  |
| 4 | 1.009 | 0.358 | 1.000 |  |  |  |
| 5 | 1.027 | 1.189 | 1.000 |  | $0<y[0.1]$ |  |
| 6 | 1.129 | 5.63 | 0.000 |  | $0>y$ [0.104] | I > h [0.092] |
| 7 | 1.043 | 1.869 | 1.000 |  |  |  |
| 8 | 0.926 | -3.18 | 0.083 |  |  | $1>\mathrm{h}$ [0.097] |
| 9 | 0.931 | -3.03 | 0.130 |  |  |  |
| 10 | 1.008 | 0.39 | 1.000 |  |  |  |
| 11 | 1.071 | 3.218 | 0.076 |  | $0<y[0.094]$ | $\mathrm{l}<\mathrm{h}$ [0.113] |
| 12 | 1.182 | 7.89 | 0.000 |  |  | $\mathrm{l}<\mathrm{h}$ [0.111] |
| 13 | 0.912 | -3.87 | 0.008 |  |  |  |
| 14 | 1.023 | 0.925 | 1.000 |  | $0<y[0.105]$ |  |
| 15 | 1.16 | 5.81 | 0.000 |  |  |  |
| 16 | 0.951 | -2.04 | 1.000 |  |  |  |
| 17 | 0.896 | -4.5 | 0.001 |  | o > y [0.076] | I > h [0.149] |
| 18 | 0.939 | -2.84 | 0.228 |  | o > y [0.074] |  |
| 19 | 0.862 | -6.43 | 0.000 |  |  |  |
| 20 | 0.802 | -9.22 | 0.000 |  |  | $1>\mathrm{h}$ [0.095] |
| 21 | 0.905 | -4.01 | 0.005 |  |  | $1>\mathrm{h}$ [0.115] |
| 22 | 0.911 | -3.92 | 0.006 |  |  | $1>\mathrm{h}$ [0.075] |
| 23 | 0.895 | -4.57 | 0.000 |  |  | $1>\mathrm{h}$ [0.079] |
| 24 | 0.942 | -2.46 | 0.662 |  |  |  |
| 25 | 1.022 | 0.982 | 1.000 |  | $0<y[0.167]$ | $\mathrm{l}<\mathrm{h}$ [0.141] |
| 26 | 1 | -0.01 | 1.000 | $\mathrm{f}<\mathrm{m}$ [0.083] | $0<y[0.131]$ |  |
| 27 | 0.955 | -2.05 | 1.000 | $\mathrm{f}<\mathrm{m}$ [0.091] | $0<y[0.157]$ |  |
| 28 | 1.103 | 4.6 | 0.000 |  |  | I<h [0.143] |
| 29 | 1.143 | 6.2 | 0.000 |  | $0<y[0.196]$ |  |
| 30 | 1.122 | 5.08 | 0.000 |  |  | l < h [0.227] |
| 31 | 1.073 | 3.28 | 0.066 |  |  | $1>\mathrm{h}$ [0.108] |
| 32 | 0.816 | -8.26 | 0.000 |  | o > y [0.099] | $1>\mathrm{h}$ [0.173] |
| 33 | 0.887 | -5.11 | 0.000 |  | o > y [0.085] | I<h [0.091] |
| 34 | 1.121 | 5.4 | 0.000 |  |  | I<h [0.095] |
| 35 | 1.229 | 10 | 0.000 |  |  | $1<\mathrm{h}$ [0.071] |
| 36 | 1.018 | 0.778 | 1.000 | $\mathrm{f}>\mathrm{m}$ [0.093] |  | $1>\mathrm{h}$ [0.093] |
| 37 | 1.014 | 0.557 | 1.000 |  |  | I<h [0.114] |
| 38 | 1.221 | 9.74 | 0.000 |  | $0>y$ [0.161] |  |
| 39 | 0.914 | -3.83 | 0.009 |  | o > y [0.104] |  |


| 40 | 0.855 | -6.6 | $\mathbf{0 . 0 0 0}$ | $\mathrm{f}<\mathrm{m}[0.079]$ |  |  |
| ---: | ---: | ---: | ---: | :--- | :--- | :--- |
| 41 | 0.999 | -0.06 | 1.000 |  | $\mathrm{o}<\mathrm{y}[0.119]$ | $\mathrm{I}<\mathrm{h}[0.109]$ |
| 42 | 0.982 | -0.75 | 1.000 |  | $\mathrm{o}<\mathrm{y}[0.098]$ | $\mathrm{I}<\mathrm{h}[0.098]$ |
| 43 | 0.896 | -4.36 | $\mathbf{0 . 0 0 1}$ |  |  |  |
| 44 | 1.025 | 1.107 | 1.000 |  |  |  |
| 45 | 1.284 | 11.5 | $\mathbf{0 . 0 0 0}$ | $\mathrm{f}>\mathrm{m}[0.086]$ | $\mathrm{o}>\mathrm{y}[0.289]$ | $\mathrm{I}>\mathrm{h}[0.207]$ |
| 46 | 1.025 | 1.111 | 1.000 |  |  |  |
| 47 | 1.179 | 7.81 | $\mathbf{0 . 0 0 0}$ |  |  |  |

Table A2: Results of Step 3 of the development of the HLS-EU-Q12

| Country | Item | infit | $t$ | $p$ Holm | DIF [magnitude] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Gender | Age | Education |
| AT | 2 | 1.011 | 0.192 | 1 |  | o > y [0.21] |  |
|  | 4 | 1.047 | 0.765 | 1 |  |  |  |
|  | 5 | 1.014 | 0.254 | 1 |  |  |  |
|  | 7 | 1.021 | 0.374 | 1 |  |  |  |
|  | 8 | 0.959 | -0.66 | 1 |  |  | I > h [0.217] |
|  | 10 | 0.987 | -0.2 | 1 |  |  |  |
|  | 11 | 1.096 | 1.641 | 1 |  |  | I < h [0.191] |
|  | 13 | 0.934 | -1.1 | 1 |  |  |  |
|  | 16 | 0.962 | -0.61 | 1 |  |  |  |
|  | 18 | 0.876 | -2.18 | 1 |  |  |  |
|  | 21 | 0.968 | -0.48 | 1 |  |  |  |
|  | 23 | 0.973 | -0.39 | 1 |  |  |  |
|  | 24 | 0.953 | -0.77 | 1 |  |  |  |
|  | 28 | 1.113 | 1.915 | 1 |  |  |  |
|  | 31 | 1.085 | 1.448 | 1 |  |  | I < h [0.203] |
|  | 33 | 0.898 | -1.78 | 1 |  |  |  |
|  | 36 | 1.114 | 1.828 | 1 |  |  |  |
|  | 37 | 1.119 | 1.919 | 1 |  |  |  |
|  | 39 | 0.917 | -1.43 | 1 |  |  |  |
|  | 42 | 1.001 | 0.027 | 1 |  |  |  |
|  | 43 | 0.94 | -0.99 | 1 |  |  |  |
|  | 44 | 1.108 | 1.813 | 1 |  |  |  |
| BG | 2 | 1.068 | 1.083 | 1 |  |  |  |
|  | 4 | 0.981 | -0.26 | 1 |  |  |  |
|  | 5 | 0.949 | -0.78 | 1 |  |  |  |
|  | 7 | 1.022 | 0.363 | 1 |  |  |  |
|  | 8 | 0.885 | -1.86 | 1 |  |  |  |
|  | 10 | 1.035 | 0.566 | 1 |  |  |  |
|  | 11 | 1.176 | 2.667 | 0.459 |  |  |  |
|  | 13 | 0.925 | -1.21 | 1 |  |  |  |
|  | 16 | 0.94 | -0.89 | 1 |  |  |  |
|  | 18 | 1.036 | 0.583 | 1 |  |  |  |
|  | 21 | 0.942 | -0.89 | 1 |  |  |  |
|  | 23 | 0.946 | -0.81 | 1 |  |  |  |
|  | 24 | 0.921 | -1.23 | 1 |  |  |  |
|  | 28 | 1.148 | 2.214 | 1 |  |  |  |
|  | 31 | 0.999 | -0 | 1 |  |  |  |
|  | 33 | 1.037 | 0.605 | 1 |  | o > y [0.286] |  |
|  | 36 | 1.164 | 2.499 | 0.686 |  |  |  |
|  | 37 | 1.053 | 0.766 | 1 |  |  |  |
|  | 39 | 0.923 | -1.23 | 1 |  |  |  |
|  | 42 | 1.008 | 0.135 | 1 |  |  |  |
|  | 43 | 0.901 | -1.55 | 1 |  |  |  |


| Country | 44 | 1.15 | 2.302 | 1 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Item | infit | $t$ | $p$ Holm | DIF [magnitude] |  |  |
|  |  |  |  |  | Gender | Age | Education |
| EL | 2 | 0.901 | -1.67 | 1 |  | o > y [0.208] |  |
|  | 4 | 0.983 | -0.24 | 1 |  |  |  |
|  | 5 | 1.019 | 0.309 | 1 |  |  |  |
|  | 7 | 1.109 | 1.713 | 1 |  |  |  |
|  | 8 | 1.053 | 0.8 | 1 |  |  |  |
|  | 10 | 1.034 | 0.588 | 1 |  |  |  |
|  | 11 | 1.011 | 0.205 | 1 |  |  |  |
|  | 13 | 0.823 | -2.94 | 0.204 |  |  |  |
|  | 16 | 1.18 | 2.589 | 0.568 |  |  |  |
|  | 18 | 0.934 | -1.08 | 1 |  |  |  |
|  | 21 | 0.901 | -1.48 | 1 |  |  |  |
|  | 23 | 1.001 | 0.031 | 1 |  | o < y [0.239] |  |
|  | 24 | 1.015 | 0.246 | 1 |  |  |  |
|  | 28 | 1.202 | 3.208 | 0.084 |  |  |  |
|  | 31 | 1.109 | 1.751 | 1 |  |  |  |
|  | 33 | 0.882 | -1.88 | 1 |  | o > y [0.246] |  |
|  | 36 | 1.002 | 0.052 | 1 |  |  |  |
|  | 37 | 1.035 | 0.518 | 1 |  |  |  |
|  | 39 | 0.86 | -2.23 | 1 |  | o > y [0.248] |  |
|  | 42 | 0.937 | -0.88 | 1 |  |  |  |
|  | 43 | 0.889 | -1.67 | 1 |  |  |  |
|  | 44 | 1.206 | 3.194 | 0.087 |  |  |  |
| ES | 2 | 1.071 | 1.001 | 1 |  | o > y [0.221] |  |
|  | 4 | 0.945 | -0.69 | 1 |  |  |  |
|  | 5 | 0.987 | -0.16 | 1 |  |  |  |
|  | 7 | 0.899 | -1.42 | 1 |  |  |  |
|  | 8 | 0.998 | -0 | 1 |  |  |  |
|  | 10 | 1.058 | 0.905 | 1 |  |  |  |
|  | 11 | 1.082 | 1.253 | 1 |  |  |  |
|  | 13 | 0.91 | -1.41 | 1 |  |  |  |
|  | 16 | 1.036 | 0.579 | 1 |  |  |  |
|  | 18 | 1.162 | 2.398 | 1 |  |  |  |
|  | 21 | 0.894 | -1.35 | 1 |  |  |  |
|  | 23 | 0.965 | -0.53 | 1 |  |  |  |
|  | 24 | 0.955 | -0.58 | 1 |  |  |  |
|  | 28 | 0.982 | -0.26 | 1 |  |  |  |
|  | 31 | 1.092 | 1.461 | 1 |  |  |  |
|  | 33 | 1.023 | 0.315 | 1 |  |  |  |
|  | 36 | 1.018 | 0.282 | 1 |  |  |  |
|  | 37 | 0.962 | -0.46 | 1 |  |  |  |
|  | 39 | 0.988 | -0.17 | 1 |  |  |  |
|  | 42 | 0.969 | -0.37 | 1 |  |  |  |
|  | 43 | 0.982 | -0.23 | 1 |  |  |  |
|  | 44 | 1.067 | 0.876 | 1 |  |  |  |


| Country | Item | infit | $t$ | $p$ Holm | DIF [magnitude] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Gender | Age | Education |
| IE | 2 | 1.134 | 1.816 | 1 |  |  |  |
|  | 4 | 1.145 | 1.846 | 1 |  |  |  |
|  | 5 | 0.99 | -0.11 | 1 |  |  |  |
|  | 7 | 1.077 | 1.142 | 1 |  |  |  |
|  | 8 | 1.003 | 0.069 | 1 |  |  |  |
|  | 10 | 0.97 | -0.46 | 1 |  |  |  |
|  | 11 | 1.057 | 0.913 | 1 |  |  |  |
|  | 13 | 0.986 | -0.17 | 1 |  |  |  |
|  | 16 | 1.032 | 0.444 | 1 |  |  |  |
|  | 18 | 1.081 | 1.233 | 1 |  |  |  |
|  | 21 | 0.923 | -1.06 | 1 |  |  |  |
|  | 23 | 0.93 | -0.93 | 1 |  |  |  |
|  | 24 | 0.893 | -1.5 | 1 |  |  |  |
|  | 28 | 1.22 | 3.242 | 0.075 |  |  |  |
|  | 31 | 1.065 | 1.018 | 1 |  |  |  |
|  | 33 | 0.897 | -1.54 | 1 |  |  |  |
|  | 36 | 1.087 | 1.131 | 1 |  |  |  |
|  | 37 | 0.974 | -0.36 | 1 |  |  |  |
|  | 39 | 0.994 | -0.06 | 1 |  |  |  |
|  | 42 | 1.045 | 0.63 | 1 |  |  |  |
|  | 43 | 0.914 | -1.2 | 1 |  |  |  |
|  | 44 | 0.946 | -0.76 | 1 |  |  |  |
| NL | 2 | 1.081 | 1.102 | 1 |  | o > y [0.231] |  |
|  | 4 | 1.039 | 0.514 | 1 |  |  |  |
|  | 5 | 1.037 | 0.553 | 1 | $f>m[0.233]$ |  |  |
|  | 7 | 0.986 | -0.19 | 1 |  |  |  |
|  | 8 | 0.909 | -1.39 | 1 |  |  |  |
|  | 10 | 0.981 | -0.29 | 1 |  |  |  |
|  | 11 | 1.082 | 1.278 | 1 |  |  |  |
|  | 13 | 1.026 | 0.399 | 1 |  |  |  |
|  | 16 | 0.906 | -1.24 | 1 |  |  |  |
|  | 18 | 0.945 | -0.8 | 1 |  |  |  |
|  | 21 | 0.892 | -1.42 | 1 |  |  | I > h [0.309] |
|  | 23 | 0.959 | -0.53 | 1 |  |  |  |
|  | 24 | 0.981 | -0.26 | 1 |  |  |  |
|  | 28 | 1.06 | 0.977 | 1 |  |  |  |
|  | 31 | 1.097 | 1.566 | 1 |  |  |  |
|  | 33 | 1.012 | 0.204 | 1 |  |  |  |
|  | 36 | 1.084 | 1.137 | 1 |  |  |  |
|  | 37 | 1.058 | 0.835 | 1 |  |  |  |
|  | 39 | 1.003 | 0.057 | 1 |  |  |  |
|  | 42 | 1.011 | 0.184 | 1 |  | o < y [0.336] |  |
|  | 43 | 0.936 | -0.88 | 1 | $\mathrm{f}<\mathrm{m}[0.226]$ |  |  |
|  | 44 | 1.176 | 2.638 | 0.467 |  |  |  |


| Country | Item | infit | $t$ | $p$ Holm | DIF [magnitude] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Gender | Age | Education |
| PL | 2 | 1.017 | 0.279 | 1 |  |  |  |
|  | 4 | 1.154 | 1.997 | 1 |  |  |  |
|  | 5 | 1.17 | 2.261 | 1 |  |  |  |
|  | 7 | 0.908 | -1.3 | 1 |  |  |  |
|  | 8 | 0.946 | -0.74 | 1 |  |  |  |
|  | 10 | 1.047 | 0.721 | 1 |  |  |  |
|  | 11 | 1.085 | 1.23 | 1 |  |  |  |
|  | 13 | 0.866 | -1.92 | 1 |  |  |  |
|  | 16 | 0.972 | -0.39 | 1 |  |  |  |
|  | 18 | 1.011 | 0.188 | 1 |  |  |  |
|  | 21 | 0.853 | -2.13 | 1 |  |  |  |
|  | 23 | 0.859 | -1.98 | 1 |  |  |  |
|  | 24 | 0.961 | -0.52 | 1 |  |  |  |
|  | 28 | 0.968 | -0.46 | 1 |  |  |  |
|  | 31 | 1.057 | 0.852 | 1 |  |  |  |
|  | 33 | 0.947 | -0.75 | 1 |  |  |  |
|  | 36 | 1.132 | 1.745 | 1 |  |  |  |
|  | 37 | 1.095 | 1.242 | 1 |  |  |  |
|  | 39 | 1.033 | 0.451 | 1 |  |  |  |
|  | 42 | 1.033 | 0.472 | 1 |  |  |  |
|  | 43 | 0.936 | -0.89 | 1 |  |  |  |
|  | 44 | 1.016 | 0.244 | 1 |  |  |  |
| DE | 2 | 1.05 | 0.814 | 1 |  |  |  |
|  | 4 | 1.021 | 0.336 | 1 |  |  |  |
|  | 5 | 0.999 | -0.01 | 1 |  |  |  |
|  | 7 | 1.079 | 1.283 | 1 |  |  |  |
|  | 8 | 0.876 | -1.8 | 1 |  |  | I > h [0.234] |
|  | 10 | 1.083 | 1.362 | 1 |  |  |  |
|  | 11 | 1.052 | 0.868 | 1 |  |  |  |
|  | 13 | 0.902 | -1.65 | 1 |  |  |  |
|  | 16 | 0.877 | -1.79 | 1 |  |  |  |
|  | 18 | 1.038 | 0.627 | 1 |  |  |  |
|  | 21 | 0.826 | -2.71 | 0.372 |  |  |  |
|  | 23 | 0.839 | -2.59 | 0.493 |  |  |  |
|  | 24 | 1.005 | 0.099 | 1 |  |  |  |
|  | 28 | 1.226 | 3.473 | 0.032 |  |  |  |
|  | 31 | 1.108 | 1.77 | 1 |  |  |  |
|  | 33 | 0.956 | -0.7 | 1 |  |  |  |
|  | 36 | 1.459 | 5.751 | 0.000 |  |  |  |
|  | 37 | 0.89 | -1.83 | 1 |  |  |  |
|  | 39 | 0.901 | -1.64 | 1 |  |  |  |
|  | 42 | 1.175 | 2.691 | 0.384 |  |  |  |
|  | 43 | 0.864 | -2.26 | 1 |  |  |  |
|  | 44 | 1.074 | 1.221 | 1 |  |  |  |

Table A3: Results of Step 4 of the development of the HLS-EU-Q12


| Country | Item | infit | t | pHolm | DIF [magnitude] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Gender | Age | Education |
| ES | 4 | 1 | -0.36 | 1 |  |  |  |
|  | 7 | 0.9 | -1.236 | 1 |  |  |  |
|  | 10 | 1.1 | 1.646 | 1 |  |  |  |
|  | 16 | 1 | -0.003 | 1 |  |  |  |
|  | 18 | 1.1 | 1.995 | 1 |  |  |  |
|  | 23 | 0.9 | -1.269 | 1 |  |  |  |
|  | 24 | 0.9 | -0.96 | 1 |  |  |  |
|  | 31 | 1.1 | 1.487 | 1 |  |  |  |
|  | 33 | 1 | 0.256 | 1 |  |  |  |
|  | 37 | 0.9 | -0.704 | 1 |  |  |  |
|  | 42 | 1 | -0.449 | 1 |  |  |  |
|  | 44 | 1 | 0.349 | 1 |  |  |  |
| IE | 4 | 1.2 | 2.017 | 1 |  |  |  |
|  | 7 | 1.1 | 1.398 | 1 |  |  |  |
|  | 10 | 1 | 0.363 | 1 |  |  |  |
|  | 16 | 1.1 | 0.747 | 1 |  |  |  |
|  | 18 | 1.1 | 1.146 | 1 |  |  |  |
|  | 23 | 0.9 | -1.067 | 1 |  |  |  |
|  | 24 | 0.9 | -1.604 | 1 |  |  |  |
|  | 31 | 1.1 | 1.516 | 1 |  |  |  |
|  | 33 | 0.9 | -1.157 | 1 |  |  |  |
|  | 37 | 1 | -0.353 | 1 |  |  |  |
|  | 42 | 1.1 | 0.793 | 1 |  |  |  |
|  | 44 | 1 | -0.593 | 1 |  |  |  |
| NL | 4 | 1 | 0.468 | 1 |  |  |  |
|  | 7 | 1 | -0.215 | 1 |  |  |  |
|  | 10 | 1 | 0.337 | 1 |  |  |  |
|  | 16 | 0.9 | -1.038 | 1 |  |  |  |
|  | 18 | 1 | -0.45 | 1 |  |  |  |
|  | 23 | 1 | -0.634 | 1 |  |  |  |
|  | 24 | 1 | -0.245 | 1 |  |  |  |
|  | 31 | 1.1 | 1.063 | 1 |  |  |  |
|  | 33 | 1 | 0.23 | 1 |  |  |  |
|  | 37 | 1.1 | 1.134 | 1 |  |  |  |
|  | 42 | 1 | -0.283 | 1 |  | o < y [0.312] |  |
|  | 44 | 1.2 | 2.453 | 0.496 |  |  |  |


|  |  |  |  |  | DIF [magnitude] |  |  |
| ---: | ---: | ---: | ---: | ---: | :--- | :--- | :--- |
| Country | Item | infit | t | pHolm | Gender | Age | Education |
| PL | 4 | 1.1 | 1.666 | 1 |  |  |  |
|  | 7 | 1 | -0.484 | 1 |  |  |  |
|  | 10 | 1.1 | 0.904 | 1 |  |  |  |
|  | 16 | 1 | -0.461 | 1 |  |  |  |
|  | 18 | 1 | -0.182 | 1 |  |  |  |
|  | 23 | 0.9 | -1.973 | 1 |  |  |  |
|  | 24 | 1 | -0.585 | 1 |  |  |  |
|  | 31 | 1.1 | 0.78 | 1 |  |  |  |
|  | 33 | 0.9 | -1.187 | 1 |  |  |  |
|  | 37 | 1.1 | 1.376 | 1 |  |  |  |
|  | 42 | 1 | 0.661 | 1 |  |  |  |
|  | 44 | 1 | -0.042 | 1 |  |  |  |
|  | 4 | 1.1 | 1.362 | 1 |  |  |  |
|  | 7 | 1.1 | 1.01 | 1 |  |  |  |
|  | 10 | 1.1 | 1.832 | 1 |  |  |  |
|  | 16 | 0.9 | -0.823 | 1 |  |  |  |
|  | 18 | 1 | 0.24 | 1 |  |  |  |
|  | 23 | 0.8 | -2.493 | 0.367 |  |  |  |
|  | 24 | 1 | -0.33 | 1 |  |  |  |
|  | 31 | 1.1 | 1.507 | 1 |  |  |  |
|  | 33 | 0.9 | -1.056 | 1 |  |  |  |
|  | 37 | 0.9 | -1.778 | 1 |  |  |  |
|  | 42 | 1.2 | 2.513 | 0.359 |  |  |  |
|  | 44 | 1 | 0.777 | 1 |  |  |  |

Table A4: Deviance and information criteria for the different questionnaire versions

|  | HLS-EU-Q47 | HLS-EU-Q16 | HLS-EU-Q12 | HLS-Q12 | HL-SF12 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Deviance |  |  |  |  |  |
| AT ( $\mathrm{n}=537$ ) | 49611.48 | 16967.97 | 13130.16 | 13392.91 | 13078.13 |
| BG ( $\mathrm{n}=509$ ) | 43225.61 | 14486.94 | 11309.28 | 11773.31 | 12191.50 |
| EL ( $\mathrm{n}=517$ ) | 44260.85 | 14889.04 | 11713.98 | 11988.47 | 12226.51 |
| ES ( $\mathrm{n}=490$ ) | 36357.60 | 12388.97 | 9370.56 | 9605.61 | 10237.56 |
| IE ( $\mathrm{n}=492$ ) | 40568.56 | 14194.64 | 10756.19 | 11111.85 | 11140.04 |
| NL ( $\mathrm{n}=498$ ) | 40929.14 | 13574.60 | 10749.74 | 10963.51 | 10759.08 |
| PL ( $\mathrm{n}=493$ ) | 36736.61 | 12738.11 | 9835.49 | 9911.04 | 10449.65 |
| DE ( $\mathrm{n}=512$ ) | 45168.99 | 15246.92 | 12001.58 | 12457.90 | 12036.25 |


| $\boldsymbol{A I C}$ |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| AT | 49895.48 | 17065.967 | 13204.16 | 13466.91 | 13152.13 |
| BG | 43509.61 | 14584.94 | 11383.28 | 11847.31 | 12265.50 |
| EL | 44544.85 | 14987.042 | 11787.98 | 12062.47 | 12300.51 |
| ES | 36641.6 | 12486.973 | 9444.56 | 9679.61 | 10311.56 |
| IE | 40852.56 | 14292.642 | 10830.19 | 11185.85 | 11214.04 |
| NL | 41213.14 | 13672.603 | 10823.74 | 11037.51 | 10833.08 |
| PL | 37020.61 | 12836.113 | 9909.49 | 9985.04 | 10523.65 |
| DE | 45452.99 | 15344.922 | 12075.58 | 12531.90 | 12110.245 |


| CAIC |  |  |  |  |  |  | 50646.09 | 17324.98 | 13399.74 | 13662.49 | 13347.71 |
| :--- | ---: | ---: | ---: | ---: | ---: | :---: | :---: | :---: | :---: | :---: | :---: |
| AT | 44252.62 | 14841.33 | 11576.88 | 12040.91 | 12459.11 |  |  |  |  |  |  |
| BG | 45290.07 | 15244.20 | 11982.16 | 12256.65 | 12494.69 |  |  |  |  |  |  |
| EL | 37379.21 | 12741.50 | 9636.75 | 9871.80 | 10503.75 |  |  |  |  |  |  |
| ES | 41590.74 | 14547.37 | 11022.54 | 11378.19 | 11406.38 |  |  |  |  |  |  |
| IE | 41953.05 | 13927.92 | 11016.53 | 11230.30 | 11025.87 |  |  |  |  |  |  |
| NL | 37759.08 | 13090.94 | 10101.91 | 10177.45 | 10716.07 |  |  |  |  |  |  |
| PL | 46196.83 | 15601.60 | 12269.40 | 12725.72 | 12304.07 |  |  |  |  |  |  |
| DE |  |  |  |  |  |  |  |  |  |  |  |


| AICc |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| AT | 49998.55 | 17076.03 | 13209.79 | 13472.55 | 13157.77 |
| BG | 43620.57 | 14595.62 | 11389.25 | 11853.28 | 12271.47 |
| EL | 44653.44 | 14997.53 | 11793.86 | 12068.34 | 12306.38 |
| ES | 36758.64 | 12498.11 | 9450.78 | 9685.832 | 10317.78 |
| IE | 40968.93 | 14303.73 | 10836.39 | 11192.04 | 11220.23 |
| NL | 31327.54 | 13683.54 | 10829.85 | 11043.62 | 10839.20 |
| PL | 45563.05 | 15355.53 | 12081.51 | 12537.83 | 12116.18 |
| DE | 50504.09 | 17275.98 | 13362.74 | 13625.49 | 13310.71 |
| BIC | 44110.62 | 14792.33 | 11539.88 | 12003.91 | 12422.11 |
| AT | 45148.07 | 15195.20 | 11945.16 | 12219.65 | 12457.69 |
| BG | 37237.21 | 12692.50 | 9599.75 | 9834.80 | 10466.75 |
| EL | 41448.74 | 14498.37 | 10985.54 | 11341.19 | 11369.38 |
| ES | 41811.05 | 13878.92 | 10979.53 | 11193.30 | 10988.87 |
| IE | 37617.08 | 13041.94 | 10064.91 | 10140.45 | 10679.07 |
| NL | 46054.83 | 15552.6 | 12232.40 | 12688.72 | 12267.07 |
| PL | 142 | 49 | 37 | 37 | 37 |
| DE |  |  |  |  |  |
| No. of <br> estimated <br> parameters |  |  |  |  |  |

Table A5: Results of Step 4 for the HLS-EU-Q12 ${ }_{32}$

| Country | Item | infit | $t$ | $p$ Holm | DIF [magnitude] |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Gender | Age | Education |
| AT | 4 | 1.048 | 0.783 | 1 |  |  |  |
|  | 7 | 1.062 | 1.074 | 1 |  |  |  |
|  | 10 | 1.04 | 0.7 | 1 |  |  |  |
|  | 16 | 0.93 | -1.15 | 1 |  |  |  |
|  | 18 | 0.876 | -2.176 | 0.975 |  |  |  |
|  | 23 | 0.946 | -0.821 | 1 |  |  |  |
|  | 24 | 0.911 | -1.51 | 1 |  |  |  |
|  | 31 | 1.136 | 2.273 | 0.784 |  |  | l < h [0.2] |
|  | 32 | 0.898 | -1.739 | 1 |  |  |  |
|  | 37 | 1.088 | 1.439 | 1 |  |  |  |
|  | 42 | 1.011 | 0.201 | 1 |  |  |  |
|  | 44 | 1.079 | 1.356 | 1 |  |  |  |
| BG | 4 | 0.973 | -0.397 | 1 |  |  |  |
|  | 7 | 1.024 | 0.389 | 1 |  |  |  |
|  | 10 | 1.089 | 1.408 | 1 |  |  |  |
|  | 16 | 0.934 | -0.984 | 1 |  |  |  |
|  | 18 | 1.073 | 1.158 | 1 |  |  |  |
|  | 23 | 0.91 | -1.381 | 1 |  |  |  |
|  | 24 | 0.91 | -1.417 | 1 |  |  |  |
|  | 31 | 0.984 | -0.245 | 1 |  | o > y [0.185] |  |
|  | 32 | 1.008 | 0.139 | 1 |  | o > y [0.252] | I > h [0.292] |
|  | 37 | 1.007 | 0.118 | 1 |  |  |  |
|  | 42 | 1.04 | 0.593 | 1 |  |  |  |
|  | 44 | 1.13 | 2.021 | 1 |  |  |  |
| EL | 4 | 0.977 | -0.331 | 1 |  |  |  |
|  | 7 | 1.081 | 1.3 | 1 |  |  |  |
|  | 10 | 1.043 | 0.739 | 1 |  |  |  |
|  | 16 | 1.127 | 1.849 | 1 |  |  |  |
|  | 18 | 0.961 | -0.639 | 1 |  |  |  |
|  | 23 | 0.934 | -0.983 | 1 |  |  |  |
|  | 24 | 0.979 | -0.299 | 1 |  |  |  |
|  | 31 | 1.083 | 1.359 | 1 |  |  |  |
|  | 32 | 0.881 | -1.848 | 1 |  | o > y [0.189] |  |
|  | 37 | 1.001 | 0.038 | 1 |  |  |  |
|  | 42 | 0.893 | -1.546 | 1 |  |  |  |
|  | 44 | 1.095 | 1.537 | 1 |  |  |  |



|  |  |  |  |  | DIF [magnitude] |  |  |
| ---: | ---: | :--- | ---: | ---: | ---: | ---: | ---: |
| Country | Item | infit | $t$ | $p$ Holm | Gender | Age | Education |
| PL | 4 | 1.127 | 1.691 | 1 |  |  |  |
|  | 7 | 0.978 | -0.287 | 1 |  |  |  |
|  | 10 | 1.086 | 1.291 | 1 |  |  |  |
|  | 16 | 0.959 | -0.589 | 1 |  |  |  |
|  | 18 | 1.018 | 0.285 | 1 |  |  |  |
|  | 23 | 0.861 | -1.967 | 1 |  |  |  |
|  | 24 | 0.948 | -0.714 | 1 |  |  |  |
|  | 31 | 1.052 | 0.774 | 1 |  |  |  |
|  | 32 | 0.847 | -2.311 | 0.624 |  |  |  |
|  | 37 | 1.091 | 1.208 | 1 |  |  |  |
|  | 42 | 1.056 | 0.798 | 1 |  |  |  |
|  | 44 | 1.013 | 0.201 | 1 |  |  |  |
|  | 4 | 1.093 | 1.38 | 1 |  |  |  |
|  | 7 | 1.062 | 1.013 | 1 |  |  |  |
|  | 10 | 1.134 | 2.158 | 0.742 |  |  |  |
|  | 16 | 0.927 | -1.028 | 1 |  |  |  |
|  | 18 | 1.04 | 0.671 | 1 |  |  |  |
|  | 23 | 0.841 | -2.566 | 0.278 |  |  |  |
|  | 24 | 0.98 | -0.284 | 1 |  |  |  |
|  | 31 | 1.114 | 1.876 | 1 |  |  |  |
|  | 32 | 0.833 | -2.769 | 0.163 |  |  |  |
|  | 37 | 0.888 | -1.873 | 1 |  |  |  |
|  | 42 | 1.179 | 2.76 | 0.163 |  |  |  |
|  | 44 | 1.057 | 0.959 | 1 |  |  |  |

Table A6: Infit statistics for single items in the total samples of the different countries

| Country | Item | Infit | $t$ | $p$ |
| :---: | :---: | :---: | :---: | :---: |
| AT |  |  |  |  |
|  | 4 | 1.029 | 0.926 | 1.000 |
|  | 7 | 1.008 | 0.279 | 1.000 |
|  | 10 | 1.068 | 2.459 | 0.334 |
|  | 16 | 1.012 | 0.420 | 1.000 |
|  | 18 | 1.015 | 0.562 | 1.000 |
|  | 23 | 0.903 | -3.346 | 0.023 |
|  | 24 | 1.011 | 0.393 | 1.000 |
|  | 31 | 1.088 | 3.027 | 0.067 |
|  | 32 | 0.949 | -1.767 | 1.000 |
|  | 37 | 1.063 | 2.166 | 0.607 |
|  | 42 | 0.976 | -0.812 | 1.000 |
|  | 44 | 0.959 | -1.467 | 1.000 |
| BE |  |  |  |  |
|  | 4 | 1.018 | 0.405 | 1.000 |
|  | 7 | 1.059 | 1.329 | 1.000 |
|  | 10 | 0.975 | -0.577 | 1.000 |
|  | 16 | 0.970 | -0.592 | 1.000 |
|  | 18 | 1.073 | 1.666 | 1.000 |
|  | 23 | 0.904 | -2.306 | 0.612 |
|  | 24 | 0.968 | -0.729 | 1.000 |
|  | 31 | 1.177 | 3.938 | 0.003 |
|  | 32 | 0.907 | -2.054 | 1.000 |
|  | 37 | 1.054 | 1.180 | 1.000 |
|  | 42 | 0.948 | -1.213 | 1.000 |
|  | 44 | 0.935 | -1.526 | 1.000 |
| CH |  |  |  |  |
|  | 4 | 1.065 | 2.103 | 0.688 |
|  | 7 | 0.941 | -1.928 | 0.862 |
|  | 10 | 1.001 | 0.043 | 1.000 |
|  | 16 | 0.973 | -0.789 | 1.000 |


| Country | Item | Infit | $t$ | $p$ |
| :---: | :---: | :---: | :---: | :---: |
|  | 18 | 0.940 | -2.185 | 0.635 |
|  | 23 | 0.975 | -0.808 | 1.000 |
|  | 24 | 1.015 | 0.504 | 1.000 |
|  | 31 | 1.150 | 5.126 | <0.001 |
|  | 32 | 0.963 | -1.205 | 1.000 |
|  | 37 | 1.108 | 3.343 | 0.027 |
|  | 42 | 0.971 | -0.979 | 1.000 |
|  | 44 | 0.954 | -1.628 | 1.000 |
| CZ |  |  |  |  |
|  | 4 | 0.925 | -1.862 | 1.000 |
|  | 7 | 0.929 | -1.788 | 1.000 |
|  | 10 | 0.997 | -0.078 | 1.000 |
|  | 16 | 1.041 | 1.020 | 1.000 |
|  | 18 | 0.994 | -0.150 | 1.000 |
|  | 23 | 0.985 | -0.363 | 1.000 |
|  | 24 | 1.044 | 1.130 | 1.000 |
|  | 31 | 0.965 | -0.921 | 1.000 |
|  | 32 | 1.052 | 1.287 | 1.000 |
|  | 37 | 1.066 | 1.571 | 1.000 |
|  | 42 | 0.966 | -0.880 | 1.000 |
|  | 44 | 1.124 | 3.278 | 0.035 |
| DE |  |  |  |  |
|  | 4 | 1.026 | 0.795 | 1.000 |
|  | 7 | 0.961 | -1.278 | 1.000 |
|  | 10 | 0.949 | -1.608 | 1.000 |
|  | 16 | 1.007 | 0.223 | 1.000 |
|  | 18 | 0.971 | -0.958 | 1.000 |
|  | 23 | 0.940 | -1.877 | 1.000 |
|  | 24 | 0.977 | -0.758 | 1.000 |
|  | 31 | 1.012 | 0.386 | 1.000 |
|  | 32 | 0.945 | -1.785 | 1.000 |
|  | 37 | 1.150 | 4.384 | <0.001 |


| Country | Item | Infit | $t$ | $p$ |
| :---: | :---: | :---: | :---: | :---: |
|  | 42 | 1.039 | 1.319 | 1.000 |
|  | 44 | 1.037 | 1.192 | 1.000 |
| DK |  |  |  |  |
|  | 4 | 1.003 | 0.121 | 1.000 |
|  | 7 | 1.036 | 1.459 | 1.000 |
|  | 10 | 0.963 | -1.602 | 1.000 |
|  | 16 | 0.926 | -3.002 | 0.070 |
|  | 18 | 1.021 | 0.893 | 1.000 |
|  | 23 | 0.975 | -1.042 | 1.000 |
|  | 24 | 0.974 | -1.066 | 1.000 |
|  | 31 | 1.048 | 2.035 | 0.838 |
|  | 32 | 0.896 | -4.454 | <0.001 |
|  | 37 | 1.073 | 2.855 | 0.103 |
|  | 42 | 1.034 | 1.393 | 1.000 |
|  | 44 | 1.096 | 4.052 | 0.002 |
| FR |  |  |  |  |
|  | 4 | 1.171 | 4.878 | <0.001 |
|  | 7 | 1.016 | 0.495 | 1.000 |
|  | 10 | 1.020 | 0.660 | 1.000 |
|  | 16 | 1.062 | 1.815 | 1.000 |
|  | 18 | 0.964 | -1.149 | 1.000 |
|  | 23 | 0.942 | -1.713 | 1.000 |
|  | 24 | 0.920 | -2.388 | 0.390 |
|  | 31 | 1.046 | 1.472 | 1.000 |
|  | 32 | 0.956 | -1.251 | 1.000 |
|  | 37 | 1.114 | 3.071 | 0.055 |
|  | 42 | 0.929 | -2.220 | 0.554 |
|  | 44 | 0.945 | -1.758 | 1.000 |
| HU |  |  |  |  |
|  | 4 | 1.190 | 3.838 | 0.004 |
|  | 7 | 1.001 | 0.027 | 1.000 |
|  | 10 | 0.957 | -0.972 | 1.000 |


| Country | Item | Infit | $t$ | $p$ |
| :---: | :---: | :---: | :---: | :---: |
|  | 16 | 0.939 | -1.162 | 1.000 |
|  | 18 | 1.116 | 2.446 | 0.419 |
|  | 23 | 0.905 | -1.815 | 1.000 |
|  | 24 | 0.940 | -1.189 | 1.000 |
|  | 31 | 1.164 | 3.647 | 0.009 |
|  | 32 | 0.894 | -2.018 | 1.000 |
|  | 37 | 0.950 | -0.911 | 1.000 |
|  | 42 | 0.938 | -1.156 | 1.000 |
|  | 44 | 0.984 | -0.336 | 1.000 |
| IE |  |  |  |  |
|  | 4 | 1.064 | 2.773 | 0.094 |
|  | 7 | 1.049 | 2.299 | 0.301 |
|  | 10 | 1.057 | 2.752 | 0.095 |
|  | 16 | 0.944 | -1.869 | 0.606 |
|  | 18 | 1.056 | 2.742 | 0.095 |
|  | 23 | 0.904 | -4.375 | <0.001 |
|  | 24 | 0.969 | -1.323 | 0.928 |
|  | 31 | 1.205 | 9.567 | <0.001 |
|  | 32 | 0.955 | -1.743 | 0.651 |
|  | 37 | 0.958 | -1.876 | 0.606 |
|  | 42 | 1.037 | 1.612 | 0.748 |
|  | 44 | 0.952 | -2.089 | 0.441 |
| IL |  |  |  |  |
|  | 4 | 1.082 | 1.903 | 1.000 |
|  | 7 | 1.009 | 0.244 | 1.000 |
|  | 10 | 1.003 | 0.090 | 1.000 |
|  | 16 | 0.984 | -0.380 | 1.000 |
|  | 18 | 1.002 | 0.052 | 1.000 |
|  | 23 | 0.931 | -1.763 | 1.000 |
|  | 24 | 1.050 | 1.249 | 1.000 |
|  | 31 | 0.980 | -0.519 | 1.000 |
|  | 32 | 1.016 | 0.389 | 1.000 |


| Country | Item | Infit | $t$ | $p$ |
| :---: | :---: | :---: | :---: | :---: |
|  | 37 | 1.039 | 0.954 | 1.000 |
|  | 42 | 1.021 | 0.544 | 1.000 |
|  | 44 | 1.018 | 0.465 | 1.000 |
| NO |  |  |  |  |
|  | 4 | 1.046 | 1.441 | 1.000 |
|  | 7 | 1.025 | 0.895 | 1.000 |
|  | 10 | 1.081 | 2.847 | 0.097 |
|  | 16 | 0.902 | -3.123 | 0.043 |
|  | 18 | 1.040 | 1.448 | 1.000 |
|  | 23 | 0.894 | -3.716 | 0.005 |
|  | 24 | 0.945 | -1.955 | 0.708 |
|  | 31 | 1.162 | 5.603 | <0.001 |
|  | 32 | 0.963 | -1.223 | 1.000 |
|  | 37 | 0.970 | -1.033 | 1.000 |
|  | 42 | 1.025 | 0.845 | 1.000 |
|  | 44 | 1.020 | 0.725 | 1.000 |
| PT |  |  |  |  |
|  | 4 | 0.952 | -0.715 | 1.000 |
|  | 7 | 1.014 | 0.215 | 1.000 |
|  | 10 | 1.032 | 0.629 | 1.000 |
|  | 16 | 0.970 | -0.427 | 1.000 |
|  | 18 | 1.066 | 1.237 | 1.000 |
|  | 23 | 0.855 | -2.328 | 0.590 |
|  | 24 | 0.863 | -2.033 | 1.000 |
|  | 31 | 0.972 | -0.487 | 1.000 |
|  | 32 | 0.806 | -2.743 | 0.213 |
|  | 37 | 0.882 | -1.489 | 1.000 |
|  | 42 | 0.906 | -1.230 | 1.000 |
|  | 44 | 1.056 | 0.834 | 1.000 |
| RU |  |  |  |  |
|  | 4 | 1.035 | 1.338 | 1.000 |
|  | 7 | 0.961 | -1.771 | 0.966 |


| Country | Item | Infit | $t$ | $p$ |
| :---: | :---: | :---: | :---: | :---: |
|  | 10 | 0.961 | -1.943 | 0.832 |
|  | 16 | 0.976 | -0.900 | 1.000 |
|  | 18 | 0.992 | -0.365 | 1.000 |
|  | 23 | 0.932 | -2.618 | 0.168 |
|  | 24 | 0.935 | -2.868 | 0.091 |
|  | 31 | 0.969 | -1.420 | 1.000 |
|  | 32 | 0.954 | -1.901 | 0.860 |
|  | 37 | 1.077 | 2.867 | 0.091 |
|  | 42 | 0.977 | -0.980 | 1.000 |
|  | 44 | 1.088 | 3.860 | 0.003 |
| SI |  |  |  |  |
|  | 4 | 1.008 | 0.294 | 1.000 |
|  | 7 | 0.896 | -3.809 | 0.002 |
|  | 10 | 1.073 | 2.756 | 0.070 |
|  | 16 | 0.908 | -3.448 | 0.008 |
|  | 18 | 0.915 | -3.277 | 0.014 |
|  | 23 | 0.995 | -0.187 | 1.000 |
|  | 24 | 0.840 | -6.158 | <0.001 |
|  | 31 | 1.350 | 12.690 | <0.001 |
|  | 32 | 0.893 | -3.975 | 0.001 |
|  | 37 | 0.991 | -0.316 | 1.000 |
|  | 42 | 0.925 | -2.705 | 0.075 |
|  | 44 | 1.134 | 4.875 | <0.001 |
| SK |  |  |  |  |
|  | 4 | 0.988 | -0.369 | 1.000 |
|  | 7 | 0.963 | -1.176 | 1.000 |
|  | 10 | 0.997 | -0.107 | 1.000 |
|  | 16 | 1.053 | 1.570 | 1.000 |
|  | 18 | 0.945 | -1.889 | 1.000 |
|  | 23 | 0.952 | -1.556 | 1.000 |
|  | 24 | 1.009 | 0.300 | 1.000 |
|  | 31 | 1.078 | 2.559 | 0.315 |


| Country | Item | Infit | $\boldsymbol{t}$ | $\boldsymbol{p}$ |
| :---: | ---: | :---: | :---: | :---: |
|  | 32 | 0.886 | -3.628 | 0.010 |
|  | 37 | 1.067 | 1.975 | 1.000 |
|  | 42 | 1.005 | 0.166 | 1.000 |
|  | 44 | 1.092 | 3.032 | 0.078 |

Table A7: Parameter of the answer categories in the total samples of the different countries

| Country | Item | Very Difficult | Difficult | Easy | Very Easy |
| :---: | :---: | :---: | :---: | :---: | :---: |
| AT |  |  |  |  |  |
|  | 4 | 0 | 0.070 | 0.927 | 3 |
|  | 7 | 0 | 0.565 | 1.262 | 3 |
|  | 10 | 0 | 0.216 | 1.084 | 3 |
|  | 16 | 0 | 0.963 | 1.423 | 3 |
|  | 18 | 0 | 0.400 | 0.960 | 3 |
|  | 23 | 0 | 0.236 | 0.892 | 3 |
|  | 24 | 0 | 0.553 | 1.130 | 3 |
|  | 31 | 0 | 0.058 | 0.765 | 3 |
|  | 32 | 0 | -0.970 | 0.124 | 3 |
|  | 37 | 0 | 0.148 | 0.552 | 3 |
|  | 42 | 0 | -0.280 | 0.527 | 3 |
|  | 44 | 0 | 0.094 | 0.744 | 3 |
| BE |  |  |  |  |  |
|  | 4 | 0 | 1.043 | 2.032 | 3 |
|  | 7 | 0 | 0.875 | 1.884 | 3 |
|  | 10 | 0 | 0.639 | 1.727 | 3 |
|  | 16 | 0 | 0.988 | 1.885 | 3 |
|  | 18 | 0 | 0.735 | 1.851 | 3 |


| Country | Item | Very Difficult | Difficult | Easy | Very Easy |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 23 | 0 | 0.745 | 1.581 | 3 |
|  | 24 | 0 | 0.963 | 2.006 | 3 |
|  | 31 | 0 | 0.368 | 1.445 | 3 |
|  | 32 | 0 | 1.196 | 2.110 | 3 |
|  | 37 | 0 | 0.836 | 1.756 | 3 |
|  | 42 | 0 | 0.768 | 1.628 | 3 |
|  | 44 | 0 | 0.694 | 1.520 | 3 |
| CH |  |  |  |  |  |
|  | 4 | 0 | 0.518 | 1.576 | 3 |
|  | 7 | 0 | 0.646 | 1.531 | 3 |
|  | 10 | 0 | 0.404 | 1.132 | 3 |
|  | 16 | 0 | -0.588 | 0.721 | 3 |
|  | 18 | 0 | 0.452 | 1.263 | 3 |
|  | 23 | 0 | 0.272 | 1.084 | 3 |
|  | 24 | 0 | 0.199 | 1.368 | 3 |
|  | 31 | 0 | 0.108 | 0.875 | 3 |
|  | 32 | 0 | 0.687 | 1.618 | 3 |
|  | 37 | 0 | 0.278 | 0.910 | 3 |
|  | 42 | 0 | 0.458 | 1.273 | 3 |
|  | 44 | 0 | 0.489 | 1.281 | 3 |
| CZ |  |  |  |  |  |
|  | 4 | 0 | 0.499 | 1.482 | 3 |
|  | 7 | 0 | 0.617 | 1.534 | 3 |
|  | 10 | 0 | 0.486 | 1.261 | 3 |
|  | 16 | 0 | 0.594 | 1.391 | 3 |
|  | 18 | 0 | 0.755 | 1.493 | 3 |


| Country | Item | Very Difficult | Difficult | Easy | Very Easy |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 23 | 0 | 0.808 | 1.675 | 3 |
|  | 24 | 0 | 0.755 | 1.756 | 3 |
|  | 31 | 0 | 0.349 | 1.321 | 3 |
|  | 32 | 0 | 0.361 | 1.662 | 3 |
|  | 37 | 0 | 0.562 | 1.389 | 3 |
|  | 42 | 0 | 0.457 | 1.556 | 3 |
|  | 44 | 0 | 0.463 | 1.358 | 3 |
| DE |  |  |  |  |  |
|  | 4 | 0 | -0.471 | 1.052 | 3 |
|  | 7 | 0 | 0.565 | 1.549 | 3 |
|  | 10 | 0 | 0.563 | 1.349 | 3 |
|  | 16 | 0 | 0.090 | 1.402 | 3 |
|  | 18 | 0 | 0.556 | 1.570 | 3 |
|  | 23 | 0 | -0.099 | 1.134 | 3 |
|  | 24 | 0 | 0.411 | 1.604 | 3 |
|  | 31 | 0 | 0.485 | 1.421 | 3 |
|  | 32 | 0 | 0.463 | 1.825 | 3 |
|  | 37 | 0 | -0.270 | 0.647 | 3 |
|  | 42 | 0 | 0.479 | 1.419 | 3 |
|  | 44 | 0 | 0.289 | 1.328 | 3 |
| DK |  |  |  |  |  |
|  | 4 | 0 | 0.499 | 1.469 | 3 |
|  | 7 | 0 | 0.926 | 1.744 | 3 |
|  | 10 | 0 | 0.623 | 1.366 | 3 |
|  | 16 | 0 | 0.536 | 1.387 | 3 |
|  | 18 | 0 | 0.431 | 1.167 | 3 |


| Country | Item | Very Difficult | Difficult | Easy | Very Easy |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 23 | 0 | 0.562 | 1.434 | 3 |
|  | 24 | 0 | 0.765 | 1.522 | 3 |
|  | 31 | 0 | 0.574 | 1.355 | 3 |
|  | 32 | 0 | 0.746 | 1.566 | 3 |
|  | 37 | 0 | 0.372 | 0.935 | 3 |
|  | 42 | 0 | 0.599 | 1.257 | 3 |
|  | 44 | 0 | 0.759 | 1.314 | 3 |
| FR |  |  |  |  |  |
|  | 4 | 0 | -0.507 | 0.661 | 3 |
|  | 7 | 0 | -0.342 | 0.811 | 3 |
|  | 10 | 0 | 0.195 | 1.021 | 3 |
|  | 16 | 0 | -2.912 | -0.783 | 3 |
|  | 18 | 0 | 0.237 | 1.129 | 3 |
|  | 23 | 0 | -0.116 | 0.846 | 3 |
|  | 24 | 0 | 0.182 | 1.145 | 3 |
|  | 31 | 0 | 0.224 | 0.967 | 3 |
|  | 32 | 0 | 0.164 | 1.320 | 3 |
|  | 37 | 0 | -0.210 | 0.657 | 3 |
|  | 42 | 0 | 0.093 | 0.972 | 3 |
|  | 44 | 0 | 0.143 | 0.946 | 3 |
| HU |  |  |  |  |  |
|  | 4 | 0 | -0.077 | 1.154 | 3 |
|  | 7 | 0 | 0.194 | 1.178 | 3 |
|  | 10 | 0 | 0.382 | 1.015 | 3 |
|  | 16 | 0 | -0.219 | 0.921 | 3 |
|  | 18 | 0 | -0.282 | 0.377 | 3 |


| Country | Item | Very Difficult | Difficult | Easy | Very Easy |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 23 | 0 | 0.057 | 1.222 | 3 |
|  | 24 | 0 | 0.605 | 1.365 | 3 |
|  | 31 | 0 | 0.192 | 0.689 | 3 |
|  | 32 | 0 | 0.483 | 1.594 | 3 |
|  | 37 | 0 | 0.702 | 1.413 | 3 |
|  | 42 | 0 | 0.445 | 1.469 | 3 |
|  | 44 | 0 | 0.385 | 0.945 | 3 |
| IE |  |  |  |  |  |
|  | 4 | 0 | -0.423 | 0.352 | 3 |
|  | 7 | 0 | -0.533 | 0.336 | 3 |
|  | 10 | 0 | -0.047 | 1.144 | 3 |
|  | 16 | 0 | 0.679 | 0.690 | 3 |
|  | 18 | 0 | 0.277 | 0.976 | 3 |
|  | 23 | 0 | -0.239 | 0.601 | 3 |
|  | 24 | 0 | 0.169 | 0.894 | 3 |
|  | 31 | 0 | -0.175 | 0.729 | 3 |
|  | 32 | 0 | -0.953 | -0.076 | 3 |
|  | 37 | 0 | -0.309 | 0.307 | 3 |
|  | 42 | 0 | -0.387 | 0.372 | 3 |
|  | 44 | 0 | -0.129 | 0.440 | 3 |
| IL |  |  |  |  |  |
|  | 4 | 0 | 0.061 | 1.092 | 3 |
|  | 7 | 0 | 0.209 | 1.082 | 3 |
|  | 10 | 0 | 0.376 | 1.195 | 3 |
|  | 16 | 0 | -0.643 | 0.482 | 3 |
|  | 18 | 0 | 0.254 | 1.022 | 3 |


| Country | Item | Very Difficult | Difficult | Easy | Very Easy |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 23 | 0 | 0.022 | 0.791 | 3 |
|  | 24 | 0 | 0.101 | 0.840 | 3 |
|  | 31 | 0 | 0.135 | 0.788 | 3 |
|  | 32 | 0 | 0.476 | 1.308 | 3 |
|  | 37 | 0 | -0.084 | 0.801 | 3 |
|  | 42 | 0 | 0.200 | 0.710 | 3 |
|  | 44 | 0 | 0.132 | 0.735 | 3 |
| NO |  |  |  |  |  |
|  | 4 | 0 | 0.417 | 0.906 | 3 |
|  | 7 | 0 | 0.412 | 1.204 | 3 |
|  | 10 | 0 | -0.299 | 0.828 | 3 |
|  | 16 | 0 | 0.714 | 1.411 | 3 |
|  | 18 | 0 | -0.056 | 0.853 | 3 |
|  | 23 | 0 | 0.660 | 1.313 | 3 |
|  | 24 | 0 | 0.482 | 1.262 | 3 |
|  | 31 | 0 | -0.059 | 0.683 | 3 |
|  | 32 | 0 | 0.576 | 1.026 | 3 |
|  | 37 | 0 | 0.323 | 0.984 | 3 |
|  | 42 | 0 | 0.168 | 0.810 | 3 |
|  | 44 | 0 | 0.296 | 0.887 | 3 |
| PT |  |  |  |  |  |
|  | 4 | 0 | -9.311 | -22.470 | 3 |
|  | 7 | 0 | -7.554 | -16.627 | 3 |
|  | 10 | 0 | -0.922 | -1.946 | 3 |
|  | 16 | 0 | -3.440 | -33.517 | 3 |
|  | 18 | 0 | -2.439 | -4.979 | 3 |


| Country | Item | Very Difficult | Difficult | Easy | Very Easy |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 23 | 0 | -4.092 | -7.742 | 3 |
|  | 24 | 0 | -10.045 | -20.545 | 3 |
|  | 31 | 0 | -1.229 | -2.786 | 3 |
|  | 32 | 0 | -8.535 | -29.424 | 3 |
|  | 37 | 0 | -8.274 | -16.101 | 3 |
|  | 42 | 0 | -4.084 | -11.217 | 3 |
|  | 44 | 0 | -10.745 | -19.270 | 3 |
| RU |  |  |  |  |  |
|  | 4 | 0 | 0.514 | 1.544 | 3 |
|  | 7 | 0 | 0.702 | 1.555 | 3 |
|  | 10 | 0 | 0.642 | 1.438 | 3 |
|  | 16 | 0 | 0.697 | 1.667 | 3 |
|  | 18 | 0 | 0.536 | 1.480 | 3 |
|  | 23 | 0 | 0.553 | 1.610 | 3 |
|  | 24 | 0 | 0.626 | 1.639 | 3 |
|  | 31 | 0 | 0.407 | 1.310 | 3 |
|  | 32 | 0 | 0.624 | 1.899 | 3 |
|  | 37 | 0 | 0.437 | 1.330 | 3 |
|  | 42 | 0 | 0.490 | 1.438 | 3 |
|  | 44 | 0 | 0.104 | 1.101 | 3 |
| SI |  |  |  |  |  |
|  | 4 | 0 | 0.644 | 1.436 | 3 |
|  | 7 | 0 | 0.506 | 1.393 | 3 |
|  | 10 | 0 | 0.733 | 1.389 | 3 |
|  | 16 | 0 | 0.349 | 1.389 | 3 |
|  | 18 | 0 | 0.459 | 1.152 | 3 |


| Country | Item | Very Difficult | Difficult | Easy | Very Easy |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 23 | 0 | 0.512 | 1.171 | 3 |
|  | 24 | 0 | 0.703 | 1.424 | 3 |
|  | 31 | 0 | 0.362 | 0.800 | 3 |
|  | 32 | 0 | 0.695 | 1.643 | 3 |
|  | 37 | 0 | 0.444 | 1.098 | 3 |
|  | 42 | 0 | 0.523 | 1.345 | 3 |
|  | 44 | 0 | 0.682 | 1.159 | 3 |
| SK |  |  |  |  |  |
|  | 4 | 0 | 0.635 | 1.648 | 3 |
|  | 7 | 0 | 0.935 | 1.878 | 3 |
|  | 10 | 0 | 0.771 | 1.571 | 3 |
|  | 16 | 0 | 0.612 | 1.445 | 3 |
|  | 18 | 0 | 0.730 | 1.694 | 3 |
|  | 23 | 0 | 0.746 | 1.493 | 3 |
|  | 24 | 0 | 0.620 | 1.529 | 3 |
|  | 31 | 0 | 0.514 | 1.290 | 3 |
|  | 32 | 0 | 0.584 | 1.730 | 3 |
|  | 37 | 0 | 0.700 | 1.416 | 3 |
|  | 42 | 0 | 0.867 | 1.550 | 3 |
|  | 44 | 0 | 0.576 | 1.475 | 3 |

Figure A1: Frequency distributions for the response categories in the different countries









■HU.X0 ■HU.X1 ■HU.X2 ■HU.X3









[^0]:    ${ }^{1}$ The NVS is a 6-item screening instrument for functional health literacy and is based on the ability to read, understand and apply information from a nutrition label.

[^1]:    ${ }^{2}$ In the analysis including both survey types
    ${ }^{3}$ In all survey modes
    ${ }^{4}$ In the analysis including both survey types and for CAWI

[^2]:    ${ }^{5}$ The fact, that some tests for DIF are not significant for CATI in CH are partly due to the low sample size.

